

Overestimates for the Gain of Multiple Linear Approximations

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Abstract

We show that Corollary 1 of “On Multiple Linear Approximations” (Crypto 2004 – LNCS 3152) is incorrect. In particular, the value given for the gain by Corollary 1 is likely to be a significant overestimate of this quantity. Thus any data requirements for linear cryptanalysis with multiple linear approximations based on this value for the gain are highly questionable.

1 Introduction

Linear cryptanalysis [4] of a block cipher in its basic form uses a linear approximation of the form

$$\alpha^T \begin{pmatrix} \mathbf{p} \\ \mathbf{c} \end{pmatrix} = k \text{ with probability } \frac{1}{2}(1 + \epsilon),$$

where α is a data mask, k is one bit of key information, \mathbf{p} is a plaintext and \mathbf{c} is a corresponding ciphertext. The value ϵ is known as the *imbalance* or *correlation* (twice the *bias*) of the linear approximation. If $\epsilon \neq 0$, then it is possible to estimate the key bit k reasonably accurately if the number N of plaintext-ciphertext pairs is at least ϵ^{-2} [4].

Enhanced forms of linear cryptanalysis [2, 3] use a collection of m such linear approximations. Such a situation with multiple linear approximations is also considered by [1], where the *gain* of such a linear cryptanalysis is defined. The gain is an attempt to quantify the advantage of a such a linear cryptanalysis over exhaustive search.

This paper is concerned with the values given for the gain by [1]. In particular, we show that the value for the gain given there by *Corollary 1* generally greatly exceeds the value for the gain given there by *Theorem 1*.

2 Multiple Linear Approximations

We consider a linear cryptanalysis based on N plaintext-ciphertext pairs. We suppose that we have m linear approximations

$$\alpha_i^T \begin{pmatrix} \mathbf{p} \\ \mathbf{c} \end{pmatrix} = k_i \text{ with probability } \frac{1}{2}(1 + \epsilon_i)$$

for distinct data masks α_i , individual bits of key information k_i and imbalances ϵ_i ($i = 1, \dots, m$). The *capacity* \bar{c}^2 of this collection of linear approximations is given by Definition 2 of [1] to be $\bar{c}^2 = \sum_{i=1}^m \epsilon_i^2$.

For simplicity, we suppose that the m key bits k_1, \dots, k_m give m bits of information about the block cipher key. We let $\mathbf{z} = (k_1, \dots, k_m)^T$ denote the *key class*, and we denote the set of all key classes by \mathcal{Z} , so $\mathcal{Z} = \mathbb{Z}_2^m$ and $|\mathcal{Z}| = 2^m$. We let \mathbf{z}^* denote the key class containing the true key, and, without loss of generality, we suppose that $\mathbf{z}^* = 0$. We let $\mathcal{Z}^* = \mathcal{Z} \setminus \{\mathbf{z}^*\} = \mathbb{Z}_2^m \setminus \{0\}$ denote

the set of key classes not containing the true key, so $|\mathcal{Z}^*| = 2^m - 1$. We denote the m -dimensional *imbalance vector* corresponding to key class \mathbf{z} by \mathbf{c}_z , so

$$\mathbf{c}_z = ((-1)^{z_1} \epsilon_1, \dots, (-1)^{z_m} \epsilon_m)^T.$$

We note that the squared distance from such an imbalance vector to the imbalance vector for the true key class is given by

$$|\mathbf{c}_z - \mathbf{c}_{z^*}|^2 = |\mathbf{c}_z - \mathbf{c}_0|^2 = \left| -2(z_1 \epsilon_1, \dots, z_m \epsilon_m)^T \right|^2 = 4 \sum_{i=1}^m z_i^2 \epsilon_i^2.$$

3 Mathematical Concepts used to Define Gain Values

The values given for the gain by [1] can be expressed in terms of two functions, g and H_m , and a random variable X , which we now define.

The function g on the positive real numbers is defined by

$$g(x) = \phi\left(-\frac{1}{2} N^{\frac{1}{2}} x^{\frac{1}{2}}\right),$$

where ϕ denotes the cumulative distribution function for a standard normal $N(0; 1)$ random variable. We note that $g(x)$ is a convex function of x for $x > 0$ as

$$g''(x) = \frac{1}{32} \frac{1}{\sqrt{2\pi}} N^{\frac{1}{2}} e^{-\frac{Nx}{8}} (N + 4x^{-1}) x^{-\frac{1}{2}} > 0 \text{ for } x > 0.$$

The function H_m on the positive real numbers is defined by

$$H_m(x) = -\log_2(2(1 - 2^{-m})x + 2^{-m}) = -\log_2\left[2\frac{|Z^*|}{|Z|}x + \frac{1}{|Z|}\right].$$

We note that $H_m(x)$ is a decreasing function of x for $x > 0$ as

$$H'_m(x) = -\frac{1}{\log 2} \left(\frac{2(1 - 2^{-m})}{2(1 - 2^{-m})x + 2^{-m}} \right) < 0 \text{ for } x > 0.$$

The random variable X is defined by

$$X = |\mathbf{c}_z - \mathbf{c}_{z^*}|^2 \text{ with probability } |Z^*|^{-1} = (2^m - 1)^{-1} \text{ for } \mathbf{z} \in \mathcal{Z}^*.$$

Thus X is the random variable giving the squared distance of an imbalance vector for an incorrect key class from the imbalance vector for the true key class.

4 Comparison of Values for the Gain

We now compare the two values given for the gain in *Theorem 1* and *Corollary 1* of [1]. We show in Appendix A that the value γ for the gain given by *Theorem 1* is given by

$$\gamma = H_m(\mathbf{E}[g(X)]).$$

We show in Appendix B that the value $\tilde{\gamma}$ for the gain given by *Corollary 1* is given by

$$\tilde{\gamma} = H_m(g(2\bar{c}^2)) = H_m(g((1 - 2^{-m})\mathbf{E}[X])).$$

However, this value $\tilde{\gamma}$ for the gain can be well approximated by $\hat{\gamma}$, where

$$\hat{\gamma} = H_m(g(\mathbf{E}[X])).$$

We now use Jensen's inequality [6] to compare γ and $\hat{\gamma}$. As g is a convex function of the positive real numbers, Jensen's inequality shows that

$$g(\mathbf{E}[X]) \leq \mathbf{E}[g(X)].$$

Furthermore H_m is a decreasing function of the positive real numbers, so

$$\hat{\gamma} = H_m(g(\mathbf{E}[X])) \geq H_m(\mathbf{E}[g(X)]) = \gamma.$$

However, $\tilde{\gamma}$ is usually extremely well-approximated by $\hat{\gamma}$, so giving Lemma 1.

Lemma 1. The value $\tilde{\gamma}$ for the gain given by *Corollary 1* generally exceeds the value γ given for the gain by *Theorem 1*.

5 Example Values for the Gain

The important issue in the use of *Corollary 1* of [1] to give the gain is whether the overestimate of γ by $\tilde{\gamma}$ referred to in Lemma 1 gives a significant error in the value of the gain. We show by giving an example that it is indeed generally the case that the use of $\tilde{\gamma}$ given in *Corollary 1* gives a large overestimate of the gain γ given by *Theorem 1*.

For simplicity, we assume that all m linear approximations have the same imbalance ϵ , that is $\epsilon_1 = \dots = \epsilon_m = \epsilon$. The capacity of such a collection of linear approximations is clearly $\bar{c}^2 = m\epsilon^2$. In this situation, using the result given in Section 2, we have

$$|\mathbf{c}_{\mathbf{z}} - \mathbf{c}_{\mathbf{z}^*}|^2 = 4\epsilon^2 \sum_{i=1}^m z_i^2 = 4\epsilon^2 |\mathbf{z}|^2.$$

As there are $\binom{m}{l}$ such vectors $\mathbf{z} \in \mathbb{Z}_2^m$ with $|\mathbf{z}|^2 = l$, the random variable X is given by

$$X = 4\epsilon^2 l \text{ with probability } \binom{m}{l} (2^m - 1)^{-1} \quad [l = 1, \dots, m].$$

Thus X is a multiple of a censored Bin($m, \frac{1}{2}$) random variable with 0 removed, so the mean of X is given by $\mathbf{E}[X] = 4\epsilon^2 \frac{m}{2} \frac{2^m}{2^m - 1}$. We therefore obtain $g(\mathbf{E}[X])$, used to define $\hat{\gamma}$, as

$$\begin{aligned} g(\mathbf{E}[X]) &= \phi \left(-\frac{1}{2} N^{\frac{1}{2}} \epsilon \left(2m \frac{2^m}{2^m - 1} \right)^{\frac{1}{2}} \right) \\ &= \phi \left(- (N\epsilon^2)^{\frac{1}{2}} \left(\frac{m}{2} \right)^{\frac{1}{2}} \left(\frac{2^m}{2^m - 1} \right)^{\frac{1}{2}} \right). \end{aligned}$$

By contrast, the mean of $g(X)$ is given by

$$\begin{aligned}\mathbf{E}[g(X)] &= \frac{1}{2^m - 1} \sum_{l=1}^m \binom{m}{l} \phi\left(-\frac{1}{2} N^{\frac{1}{2}} 2\epsilon l^{\frac{1}{2}}\right) \\ &= \frac{1}{2^m - 1} \sum_{l=1}^m \binom{m}{l} \phi\left(- (N\epsilon^2)^{\frac{1}{2}} l^{\frac{1}{2}}\right).\end{aligned}$$

We now consider the values of these two expressions for a particular example. We suppose that there are $m = 8$ linear approximations, so the capacity $\bar{c}^2 = 8\epsilon^2$. We further suppose that we have $N = 2\epsilon^{-2}$ plaintext-ciphertext pairs, so $N\epsilon^2 = 2$. In this case we have

$$g(\mathbf{E}[X]) = 0.0023, \text{ whereas } \mathbf{E}[g(X)] = 0.0074.$$

For this example, we have $g(\mathbf{E}[X]) > 3\mathbf{E}[g(X)]$, so illustrating Jensen's inequality. However, despite Jensen's inequality, it is essentially asserted by the "proof" of *Corollary 1* that $g(\mathbf{E}[X]) = \mathbf{E}[g(X)]$. (We note that the function $f(x) = g(-x)$ is erroneously used in this "proof" instead of $g(x)$.) This example shows that this assertion, which is the basis of the "proof" of *Corollary 1*, is simply wrong. As with the discussion by [1] of probabilities for dependent data masks [5], the given "proof" of *Corollary 1* by [1] is not correct.

We now calculate the various values given for the gain in this situation, so

$$\gamma = H_8(g(\mathbf{E}[X])) = 5.75 \text{ and } \hat{\gamma} = H_8(\mathbf{E}[g(X)]) = 6.88.$$

Furthermore, a direct calculation gives $\tilde{\gamma} = 6.87$, so $\hat{\gamma}$ is obviously a very good approximation of $\tilde{\gamma}$. In this situation, *Corollary 1* overestimates the gain as given by *Theorem 1* by over one bit in six.

6 Conclusions

We have shown that the value for the gain given by *Corollary 1* of [1] is not reliable, and is in general a large overestimate of the value of the gain given by *Theorem 1*. Furthermore, the "proof" given of *Corollary 1* simply ignores Jensen's inequality, a fundamental result in probability and theoretical statistics. Any result based on this value for the gain given by *Corollary 1*, such as the theoretical data requirements for such a linear cryptanalysis, is therefore highly questionable.

References

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A Value for the Gain given by *Theorem 1*

The value γ given for the gain given by *Theorem 1* of [1] is

$$\gamma = -\log_2 \left[2 \frac{1}{|\mathcal{Z}|} \sum_{\mathbf{z} \in \mathcal{Z}^*} \phi \left(-\frac{1}{2} N^{\frac{1}{2}} |\mathbf{c}_{\mathbf{z}} - \mathbf{c}_{\mathbf{z}^*}| \right) + \frac{1}{|\mathcal{Z}|} \right].$$

Thus we have

$$\gamma = -\log_2 \left[2 (1 - 2^{-m}) \frac{1}{2^m - 1} \sum_{\mathbf{z} \neq 0} g (|\mathbf{c}_{\mathbf{z}} - \mathbf{c}_0|^2) + 2^{-m} \right].$$

However, the mean value of $g(X)$ is given by

$$\mathbf{E}[g(X)] = \frac{1}{2^m - 1} \sum_{\mathbf{z} \neq 0} g (|\mathbf{c}_{\mathbf{z}} - \mathbf{c}_0|^2),$$

so we have shown that the value γ given for the gain by *Theorem 1* is given by

$$\gamma = -\log_2 [2 (1 - 2^{-m}) \mathbf{E}[g(X)] + 2^{-m}] = H_m(\mathbf{E}[g(X)]).$$

B Value for the Gain given by *Corollary 1*

The value $\tilde{\gamma}$ given for the gain by *Corollary 1* of [1] is

$$\begin{aligned} \tilde{\gamma} &= -\log_2 \left[2 \frac{|\mathcal{Z}|-1}{|\mathcal{Z}|} \phi \left(-\left(\frac{1}{2} N \bar{c}^2 \right)^{\frac{1}{2}} \right) + \frac{1}{|\mathcal{Z}|} \right] \\ &= H_m \left(\phi \left(-\frac{1}{2} (2N \bar{c}^2)^{\frac{1}{2}} \right) \right) = H_m (g (2\bar{c}^2)). \end{aligned}$$

We express this quantity in terms of the mean of X , which is given by

$$\mathbf{E}[X] = \frac{1}{|\mathcal{Z}^*|} \sum_{\mathbf{z} \neq \mathbf{z}^*} |\mathbf{c}_{\mathbf{z}} - \mathbf{c}_{\mathbf{z}^*}|^2 = (2^m - 1)^{-1} \sum_{\mathbf{z} \neq 0} |\mathbf{c}_{\mathbf{z}} - \mathbf{c}_0|^2.$$

However, $|\mathbf{c}_{\mathbf{z}} - \mathbf{c}_0|^2 = 4 \sum_{i=1}^m z_i^2 \epsilon_i^2$ (Section 1), so we have

$$\mathbf{E}[X] = (2^m - 1)^{-1} \sum_{\mathbf{z} \neq 0} \sum_{i=1}^m 4z_i^2 \epsilon_i^2 = \frac{4}{2^m - 1} \sum_{i=1}^m \epsilon_i^2 \sum_{\mathbf{z} \neq 0} z_i^2 = \frac{4 \cdot 2^{m-1}}{2^m - 1} \bar{c}^2,$$

as the two summations in the above expression can be evaluated to give

$$\sum_{\mathbf{z} \neq 0} z_i^2 = \sum_{\mathbf{z} \neq 0} z_i = 2^{m-1} \quad \text{and} \quad \sum_{i=1}^m \epsilon_i^2 = \bar{c}^2.$$

Thus we can give the capacity in terms of the mean of X as

$$\bar{c}^2 = \frac{1}{2} (1 - 2^{-m}) \mathbf{E}[X],$$

so we can obtain

$$\phi \left(- \left(\frac{1}{2} N \bar{c}^2 \right)^{\frac{1}{2}} \right) = \phi \left(- \frac{1}{2} N^{\frac{1}{2}} \left((1 - 2^{-m}) \mathbf{E}[X] \right)^{\frac{1}{2}} \right) = g \left((1 - 2^{-m}) \mathbf{E}[X] \right).$$

This means we can express the value $\tilde{\gamma}$ given for the gain by *Corollary 1* as

$$\tilde{\gamma} = H_m \left(g \left((1 - 2^{-m}) \mathbf{E}[X] \right) \right).$$

If we now define the value

$$\hat{\gamma} = H_m \left(g \left(\mathbf{E}[X] \right) \right),$$

then clearly $\hat{\gamma}$ is a very good approximation of $\tilde{\gamma}$ when m is moderately large.