INVOLUTIVE RANDOM WALKS ON TOTAL ORDERS AND THE ANTI-DIAGONAL EIGENVALUE PROPERTY

JOHN R. BRITNELL AND MARK WILDON

ABSTRACT. This paper studies a family of random walks defined on the finite ordinals using their order reversing involutions. Starting at $x \in \{0, 1, \dots, n-1\}$, an element $y \leq x$ is chosen according to a prescribed probability distribution, and the walk then steps to n-1-y. We show that under very mild assumptions these walks are irreducible, recurrent and ergodic. We then find the invariant distributions, eigenvalues and eigenvectors of a distinguished subfamily of walks whose transition matrices have the global anti-diagonal eigenvalue property studied in earlier work by Ochiai, Sasada, Shirai and Tsuboi. We prove that this subfamily of walks is characterised by their reversibility. As a corollary, we obtain the invariant distributions and rate of convergence of the random walk on the set of subsets of $\{1, \ldots, m\}$ in which steps are taken alternately to subsets and supersets, each chosen equiprobably. We then consider analogously defined random walks on the real interval [0, 1] and use techniques from the theory of self-adjoint compact operators on Hilbert spaces to prove analogues of the main results in the discrete case.

1. Introduction

In this paper we study a family of random walks defined on the finite ordinals using their order reversing involutions: starting at $x \in \{0, 1, ..., n-1\}$, an element $y \leq x$ is chosen according to a prescribed probability distribution, and the walk then steps to n-1-y. Under very mild assumptions we show that these walks are irreducible; if irreducible, they are recurrent and ergodic with a unique invariant distribution. We then consider a distinguished subfamily of walks whose transition matrices have the global anti-diagonal eigenvalue property studied in earlier work by Ochiai, Sasada, Shirai and Tsuboi [6]. We find their invariant distributions, eigenvectors and eigenvalues and show that this subfamily is characterised by their reversibility. We also give a new proof of one of the main results of [6] which characterises the generic (in the sense of algebraic geometry) matrices having the global

Date: February 11, 2022.

 $^{2010\} Mathematics\ Subject\ Classification.$ Primary: 15B51, Secondary: 05A10, 60G10, 60J05, 15A18, 46E20.

Key words and phrases. Random walk, involution, eigenvector, eigenvalue, antidiagonal eigenvalue property, binomial transform, spectrum.

anti-diagonal eigenvalue property, and prove a related characterisation that avoids the genericity hypothesis. As a corollary of our probabilistic results we obtain the invariant distributions and rate of convergence of the random walk on the set of subsets of $\{1, \ldots, m\}$ in which steps are defined by moving to a subset and then to a superset.

In the second part we consider the analogous involutive walks on the real interval [0, 1] and use techniques from the theory of self-adjoint compact operators on Hilbert spaces to prove analogues of the main results in the discrete case. We also prove further results on a trigonometrically-weighted random walk that do not appear to have discrete analogues and consider the interplay between the discrete and continuous results.

Setup. Let [y,x] denote the interval $\{y,y+1,\ldots,x\}\subseteq\mathbb{N}_0$ and let [x] denote $\{0,1,\ldots,x\}$. Let $\star:\{0,1,\ldots,n-1\}\to\{0,1,\ldots,n-1\}$ be the order anti-involution defined by $x^\star=n-1-x$; the n will always be clear from context. It is very convenient to specify the probabilities of steps using the following definition.

Definition 1.1. A weight with domain \mathbb{N}_0 is a function γ on the set of intervals of \mathbb{N}_0 , taking values in the non-negative real numbers, such that $\gamma_{\varnothing} = 0$ and $\sum_{y \in [x]} \gamma_{[y,x]} > 0$ for all $x \in \mathbb{N}_0$. We say that a weight γ is

- strictly positive if $\gamma_{[y,x]} > 0$ for all $x, y \in n$ with $y \leqslant x$;
- atomic if $\gamma_{[y,x]} = \gamma_{[y,y]}$ for all $x, y \in n$ with $y \leqslant x$; and
- \star -symmetric if $\gamma_{[y,x]} = \gamma_{[x^{\star},y^{\star}]}$ for all $x, y \in n$.

We write γ_y for $\gamma_{[y,y]}$ and $N(\gamma)$ for the strictly positive real-valued function on \mathbb{N}_0 defined by $N(\gamma)_x = \sum_{y \in [x]} \gamma_{[y,x]}$.

For brevity, when referring to states in an involutive walk, we shall use the ordinal notation in which n is equal to the subset $\{0, 1, \ldots, n-1\}$ of \mathbb{N}_0 . Thus $x \in \{0, 1, \ldots, n-1\}$ is written as $x \in n$. We define a weight with domain n by replacing the ordinal \mathbb{N}_0 with n in Definition 1.1.

Definition 1.2. Let γ be a weight with domain containing n. The γ -weighted involutive walk on n is the Markov chain with steps defined as follows: if the current state is $x \in n$ then choose y from [x] with probability proportional to $\gamma_{[y,x]}$, and step to y^* .

Writing $P(\gamma)$ for the transition matrix of the γ -weighted involutive walk on n, Definition 1.2 asserts that

(1.1)
$$P(\gamma)_{xz} = \frac{\gamma_{[z^*,x]}}{N(\gamma)_x}.$$

for all $x, z \in n$. (Note that we number rows and columns of matrices from 0.)

General involutive walks. Our first main result gives two broadly applicable sufficient conditions for a weighted involutive walk to be ergodic.

Theorem 1.3. Let γ be a weight with domain containing n. If either

- (i) $\gamma_{[0,x]} > 0$ and $\gamma_{[x,n-1]} > 0$ for all $x \in n$; or
- (ii) $\gamma_{[x,x]} > 0$ for all $x \in n$ and $\gamma_{[x-1,x]} > 0$ for all non-zero $x \in n$ then the γ -weighted involutive walk on n is irreducible, recurrent and ergodic with a unique invariant distribution.

We are particularly concerned with reversible involutive walks.

Theorem 1.4. Let γ be a weight satisfying either condition in Theorem 1.3. The γ -weighted involutive walk on n is reversible if and only if γ factorizes as a product $\alpha\beta$ where α is an atomic weight and β is a \star -symmetric weight. Moreover, in this case the unique invariant distribution is proportional to $\alpha_{r^*}N(\alpha\beta)_r$.

Binomial weights. An important family of strictly positive weights that factorize as in Theorem 1.4 is defined for $a, b \in \mathbb{R}^{>-1}$ by

(1.2)
$$\gamma_{[y,x]}^{(a,b)} = {y+a \choose y} {b+x-y \choose x-y}.$$

(We define binomial coefficients by $\binom{r}{d} = r(r-1)\dots(r-d+1)/d!$ for all $r \in \mathbb{R}$ and $d \in \mathbb{N}_0$.) In particular, $\gamma^{(0,0)}$ is the constant weight, and $\gamma^{(1,0)}$ and $\gamma^{(0,1)}$ are the weights specifying that $y \in [x]$ is chosen with probability proportional to |[y]| = y + 1 and |[y,x]| = x - y + 1, respectively.

A natural generalization leads to two more families of weights. Observe that $\gamma_{[y,x]}^{(a,b)}$ is asymptotic to $a^yb^{x-y}/y!(x-y)!$ as $a, b \to \infty$. Setting b=ac and scaling by a^x we obtain $c^{x-y}/y!(x-y)!$. We therefore define $\gamma^{(c)}$ for $c \in \mathbb{R}^{>0}$ by

(1.3)
$$\gamma_{[y,x]}^{(c)} = \begin{pmatrix} x \\ y \end{pmatrix} c^{x-y}.$$

As a further extension, we use that $\binom{y-a'}{y} = (-1)^y \binom{a'-1}{y}$ and $\binom{x-y-b'}{x-y} = (-1)^{x-y} \binom{b'-1}{x-y}$ to extend the weights $\gamma^{(a,b)}$ to $a, b \in \mathbb{R}^{<-1}$, by defining

(1.4)
$$\delta_{[y,x]}^{(a',b')} = \binom{a'-1}{y} \binom{b'-1}{x-y}$$

for a',b'>1. If $b'\not\in\mathbb{N}_0$ then $\delta^{(a',b')}$ is a strictly positive factorizable weight with domain $\min(\lceil a'\rceil,\lceil b'\rceil)$. If $b'\in\mathbb{N}_0$ then $\delta^{(a',b')}$ is a non-negative factorizable weight with domain $\lceil a'\rceil$, satisfying hypothesis (ii) in Theorem 1.3. (The domains are chosen to be as large as possible for these positivity properties: see Lemma 7.2.) In either case, $\delta^{(a',b')}_{[y,x]}$ is equal to $(-1)^x \gamma^{(-a,-b)}_{[y,x]}$, when $\gamma^{(-a',-b')}$ is defined as a function on intervals in the obvious way.

To orient the reader we show the transition matrices $P(\gamma^{(0,0)})$, $P(\gamma^{(1,0)})$, $P(\gamma^{(0,1)})$, $P(\gamma^{(0,1)})$, $P(\gamma^{(0,1)})$, $P(\gamma^{(0,1)})$, and $P(\delta^{(0,1)})$ below for n=4. To emphasise

their characteristic *anti-triangular* structure, we use dots to denote entries that are zero because the relevant interval is empty. This convention is in force throughout.

The symmetry under swapping a and b seen by comparing $P(\gamma^{(1,0)})$, $P(\gamma^{(0,1)})$ becomes a symmetry under swapping c and 1/c, seen by comparing $P(\gamma^{(\frac{1}{2})})$ and $P(\gamma^{(2)})$. Generally, if $m \in \mathbb{N}$ and m < n then $P(\delta^{(a',m)})$ has precisely m non-zero anti-diagonal bands.

Remarkably, each matrix P above has eigenvalues $(-1)^d P_{dd^*}$ for $d \in \{0,1,2,3\}$. Up to signs, these are the entries on its anti-diagonal. This is generalized in the final claim in the theorem below.

Theorem 1.5. The $\gamma^{(a,b)}$ -, $\delta^{(a',b')}$ - and $\gamma^{(c)}$ -weighted involutive walks on any n in their domain are irreducible, recurrent, reversible and ergodic. In each case the unique invariant distribution is π where π_x is proportional to $\binom{n-1-x+a}{n-1-x}\binom{x+a+b+1}{x}$ for $\gamma^{(a,b)}$, to $\binom{n-1}{x}\binom{c+1}{c}^x$ for $\gamma^{(c)}$, and to $\binom{a'-1}{n-1-x}\binom{(a'-1)+(b'-1)}{x}$ for $\delta^{(a',b')}$. The eigenvalues of $P(\gamma^{(a,b)})$ are

$$(1.5) \qquad \qquad (-1)^d \binom{a+d}{d} / \binom{a+b+d+1}{d},$$

the eigenvalues of $P(\gamma^{(c)})$ are $(-1)^d/c^d$ and the eigenvalues of $P(\delta^{(a',b')})$ are $(-1)^d \binom{a'-1}{d} / \binom{(a'-1)+(b'-1)}{d}$, for $d \in \{0,\ldots,n-1\}$. In each case, up to the signs $(-1)^d$, these are the anti-diagonal entries of the transition matrix.

If $a, b \in \mathbb{N}_0$ then a more convenient form for the absolute values of the eigenvalues of $P(\gamma^{(a,b)})$ is $\binom{a+b+1}{a}/\binom{a+b+d+1}{b+1}$, in which d appears only once. We show after (5.6) that, as this form suggests, these eigenvalues are strictly decreasing in absolute value. Hence the second largest eigenvalue in absolute value is -(a+1)/(a+b+2). This controls the rate of convergence of the $\gamma^{(a,b)}$ -weighted involutive walk. More precisely, by Corollary 12.7 in [4],

$$\lim_{t \to \infty} \left| \left| \frac{P(\gamma^{(a,b)})_{xy}^t}{\pi_y} - 1 \right| \right|^{1/t} = \frac{a+1}{a+b+1}$$

for any $x, y \in n$, with similar results for $\gamma^{(c)}$ and $\delta^{(a',b')}$.

The anti-diagonal eigenvalue property. To motivate our second result and explain the second part of the title, we need some further notation. If an involutive walk steps from x to y^* , then we say that it makes a down-step from x to y. Given a weight γ , we define the down-step matrix $H(\gamma)$ by $H(\gamma)_{xy} = \gamma_{[y,x]}/N(\gamma)_x$. Note that $H(\gamma)$ is lower-triangular and $P(\gamma)_{xz} = H(\gamma)_{xz^*}$. Hence if J(n) is the $n \times n$ matrix defined, using Iverson bracket

notation, by $J(n)_{xy} = [x = y^*]$, having ones on its anti-diagonal and zeros elsewhere, then $P(\gamma) = H(\gamma)J(n)$. Theorem 1.5 asserts that the eigenvalues of $P(\gamma)$ are, up to signs, the eigenvalues of $H(\gamma)$, for each weight γ .

More generally, the authors of [6] say in their Definition 2.3 that a lower-triangular $n \times n$ matrix L has the anti-diagonal eigenvalue property if L is diagonalizable and the eigenvalues of LJ(n) are precisely $(-1)^d L_{dd}$ for $d \in n$. By Theorem 1.5, the down-step matrices $H(\gamma^{(a,b)})$, $H(\gamma^{(c)})$ and $H(\delta^{(a',b')})$ all have this property. The authors of [6] conclude after (2.8) that a complete classification of such matrices is infeasible. They therefore introduce the global anti-diagonal eigenvalue property, namely that for all $m \leq n$, the $m \times m$ top-left submatrix of L has the anti-diagonal property. Since the weights $\gamma^{(a,b)}$, $\gamma^{(c)}$ and $\delta^{(a',b')}$ are defined without reference to the size of the matrix, the lower-triangular down-step matrices $H(\gamma^{(a,b)})$, $H(\gamma^{(c)})$ and $H(\delta^{(a',b')})$ all have this stronger property.

One of the main results of [6] is Theorem 2.17, that a lower-triangular matrix L with eigenvalues $\lambda_0, \lambda_1, \ldots, \lambda_{n-1}$ read from top to bottom, such that $(\lambda_0, \lambda_1, \ldots, \lambda_{n-1})$ is not in an exceptional set of measure zero in \mathbb{R}^n , has the global anti-diagonal eigenvalue property if and only if L is equal to the matrix H^{λ} defined by

(1.6)
$$H^{\lambda} = B(n)\operatorname{Diag}(\lambda_0, \dots, \lambda_{n-1})B(n)^{-1}$$

where B(n) is the $n \times n$ Pascal's Triangle matrix defined by $B_{xy} = {x \choose y}$. We say that H^{λ} is a binomial transform and that the matrix $H^{\lambda}J(n)$ has the generic global diagonal eigenvalue property. The proof of Theorem 1.5 shows that $H(\gamma^{(a,b)})$, $H(\delta^{(a',b')})$, $H(\gamma^{(c)})$ are all binomial transforms. This gives another proof that the corresponding transition matrices $P(\gamma^{(a,b)})$, $P(\gamma^{(c)})$ and $P\delta^{(a',b)}$) have the generic global anti-diagonal eigenvalue property.

Since [6] appears to be unpublished, in Theorem 2.4 we give a new proof of Theorem 2.17 from [6] using some basic algebraic geometry to characterize the exceptional set as the vanishing set of a polynomial. Example 2.5 extends the examples in [6] by exhibiting a family of stochastic matrices that have the global anti-diagonal eigenvalue property, but not the generic global anti-diagonal eigenvalue property; thus the genericity assumption is essential.

As an alternative, we say that an $n \times n$ lower-triangular matrix L with diagonal entries $\lambda_0, \lambda_1, \ldots, \lambda_{n-1}$ has global anti-diagonal eigenbasis action if there is a basis of eigenvectors of L such that LJ(n) has an upper-triangular matrix in this basis with diagonal entries $\lambda_0, -\lambda_1, \ldots, (-1)^{n-1}\lambda_{n-1}$, and the same holds for all the top-left submatrices of L. Theorem 2.9 characterises matrices with global anti-diagonal eigenbasis action as precisely the binomial transforms, that is, those of the form H^{λ} . In particular, $H(\gamma^{(a,b)}), H(\delta^{(a',b')}), H(\gamma^{(c)})$ all have global anti-diagonal eigenbasis action.

Stochastic matrices with global anti-diagonal eigenbasis action. Motivated by this observation, in §6 we characterize all stochastic matrices that are binomial transforms. Define $P^{\lambda} = H^{\lambda}J(n)$.

Corollary 1.6. Let $\lambda_0, \ldots, \lambda_{n-1} \in \mathbb{R}$. The matrix P^{λ} is the transition matrix of an involutive walk if and only if $\lambda_0 = 1$ and

$$\sum_{e=0}^{z} (-1)^e {z \choose e} \lambda_{z^*+e} \geqslant 0$$

for all $z \in n$. Moreover if strict inequality holds for all $z \in n$ then $1 > \lambda_1 > \ldots > \lambda_{n-1}$, and $P_{xz}^{\lambda} > 0$ if and only if $x + z \ge n - 1$.

We extend this result in Corollary 6.4 by characterising when the involutive walk defined by P^{λ} is irreducible, recurrent and ergodic.

It is important to note that §2.2 of [6] considers the related problem of when a half-infinite lower-triangular matrix all of whose upper-left submatrices are of the form (1.6) is stochastic. The main result is Theorem 2.34 which characterises such matrices as those whose diagonal entries $\lambda_0, \lambda_1, \ldots$ are *completely monotonic*: that is

$$\sum_{e=0}^{k} (-1)^e \binom{k}{e} \lambda_{m+e} \geqslant 0$$

for all $m, k \in \mathbb{N}_0$. The authors comment at the start of §2.2 that 'similar results hold in the finite case'. This is true, but as is clear from Corollary 1.6, only a one-parameter family of inequalities is required. Moreover, the inequalities in Corollary 1.6 are specific to each n: they may all hold for $\lambda_0, \ldots, \lambda_{n-1}$ and then fail for $\lambda_0, \ldots, \lambda_{n-1}, \lambda_n$. For instance, this is the case if λ_n is large. Thus it is not possible to obtain our results directly from [6]; instead one must adapt the proofs, which require the use of the quite deep Hausdorf Moment Theorem. Another sign that the problems are somewhat different is that there is no 'limiting case' of the involutive walk, because the ordinal $\mathbb N$ has no anti-involution.

The main result in §7 further extends Corollary 1.6. It is not considered in [6]. We require the following definition.

Definition 1.7. Let P be the transition matrix of an involutive walk. We say that P is *globally reversible* if all the top-right submatrices of P define reversible involutive walks.

Again, since the weights $\gamma^{(a,b)}$, $\gamma^{(c)}$ and $\delta^{(a',b')}$ are defined without reference to the size of the matrix, the transition matrices $P(\gamma^{(a,b)})$, $P(\delta^{(a',b')})$ and $P(\gamma^{(c)})$ are all globally reversible. Remarkably, this property characterizes them.

Theorem 1.8. Let $n \ge 3$ and suppose that P^{λ} is the transition matrix of an involutive walk in which 0 is accessible from every state. Then P^{λ} is globally reversible if and only if $P^{\lambda} = P(\gamma^{(a,b)})$ for unique $a, b \in \mathbb{R}^{>-1}$, or $P^{\lambda} = P(\gamma^{(c)})$ for a unique $c \in \mathbb{R}^{>0}$, or $P^{\lambda} = P(\delta^{(a',b')})$ for unique a', $b' \in \mathbb{R}^{>n-1}$ with either $b' \in \mathbb{R}^{>n-1}$ or $b' \in \mathbb{N}$ and $b' \ge 2$.

In fact we prove a somewhat sharper result, stated in Proposition 7.3, which implies, for instance, that if n=10 and $\lambda=(1,\frac{2}{3},\nu,\ldots)$ then P^{λ} is globally reversible if and only if P^{λ} is of the form required in Theorem 1.8 and the third largest eigenvalue ν is either in the interval $(\frac{1}{3},\frac{10}{23})$ or is one of the four exceptional values $\frac{10}{23},\frac{13}{30},\frac{22}{51}$ and $\frac{3}{7}$, corresponding to the weights $\delta^{(17,9)}$, $\delta^{(15,8)}$, $\delta^{(13,7)}$ and $\delta^{(11,6)}$, respectively. See Example 7.4 for full details.

We mention that while it is easy to find non-stochastic matrices of the form P^{λ} that satisfy the detailed balanced equations with respect to a 1-eigenvector, the only stochastic examples known to the authors have the stronger global reversibility property. We therefore make the following conjecture which we offer as an open problem.

Conjecture 1.9. An $n \times n$ stochastic matrix of the form P^{λ} is reversible if and only if P^{λ} is one of the matrices in Theorem 1.8, or $P^{\lambda} = J(n)$.

This conjecture has been verified by computer algebra for $n \leq 8$.

Continuous walks. The continuous analogue of the discrete involutive walk is defined on the interval [0,1]. Starting at a state $x \in [0,1]$, an element $y \in [0, x]$ is chosen according to a probability distribution specified by a real weight (as defined in Definition 8.1); the walk then steps to 1-y. In §8 and §9 we prove analogues of all the results in §3 and §5 in the continuous setting. We mention here that, by Theorem 9.6, the analogue of the $\gamma^{(a,b)}$ weighted involutive walk has discrete spectrum $(-1)^d \binom{a+d}{d} / \binom{a+b+d+1}{d}$ for $d \in \mathbb{N}_0$; suitably shifted Jacobi functions form the corresponding complete orthonormal basis of eigenfunctions for the analogue of left multiplication by the transition matrix. Thus the right-eigenvectors for the eigenvalues identified in Theorem 1.5 are the discrete analogue of the Jacobi functions. We believe these discrete analogoues are of considerable combinatorial interest. We identify some of their properties in §5.4 and compare then with the Jacobi functions in the final subsection §9.3. We also use the extended example in §8.3 to show that one plausible generalization of the anti-diagonal eigenvalue property to the continuous case is false: the continuous case therefore has a richer theory which deserves further study.

Involutive walks on subsets. Since $P(\gamma)^2 = H(\gamma) (J(n)H(\gamma)J(n))$, taking two steps at a time in a discrete involutive walk gives the Markov chain in which each step is a down-step followed by an up-step. An immediate corollary of Theorem 1.5 is that, when weighted by $\gamma^{(a,b)}$ (or one of its generalizations) this down-up walk on n is irreducible, reversible, ergodic and recurrent. Moreover its eigenvalues are the squares of the eigenvalues in Theorem 1.5. A similar result holds on [0,1] and for the up-down walk. Thus despite its more intricate definition, the involutive walk is the more fundamental of the two random processes. This is seen in the next corollary.

Corollary 1.10. Fix $m \in \mathbb{N}$ and let $0 . Let <math>\mathcal{P}$ be the set of subsets of $\{1, \ldots, m\}$.

- (i) The random walk on \mathcal{P} in which, starting from $X \in \mathcal{P}$, a subset $Y \subseteq X$ is chosen by putting each $x \in X$ in Y independently with probability p, and the walk then steps to $\{1,\ldots,m\}\setminus Y$, is irreducible, reversible, recurrent and ergodic with unique invariant distribution π where $\pi_X = p^{m-|X|}/(1+p)^m$ for each $X \in \mathcal{P}$. Its eigenvalues are $(-p)^e$ for $0 \leqslant e \leqslant m$, each with multiplicity $\binom{m}{e}$.
- (ii) Taking two steps at a time in the walk in (i) gives the walk on \mathcal{P} in which each step is a down-step followed by an up-step. This walk has the same invariant distribution, and its eigenvalues are p^{2e} for $0 \leq e \leq m$, again each with multiplicity $\binom{m}{e}$.

A connection with representation theory. The transition matrix $P(\gamma^{(0,0)})$, shown earlier when n=4, appears in [5, Example 7.5] in the context of a random walk on the indecomposable non-projective representations of $SL_2(\mathbb{F}_p)$ in characteristic p, as one of two blocks (after rearranging rows and columns) of the matrix in Figure 5 of [5], taking the parameter p to be (p-1)/2.

Outline. In §2 we give a self-contained proof of the characterisation (1.6) of matrices with the generic global anti-diagonal property first stated and proved in [6]. We then prove our sharper characterisation using the new idea of a global anti-diagonal eigenbasis introduced in Definition 2.7. Readers interested mainly in the applications to probability need only read as far as Proposition 2.3. In §3 we prove Theorems 1.3 and 1.4. In §4 we use Proposition 2.3 to prove Theorem 1.5 for the weights $\gamma^{(c)}$. In §5 we extend the method of §4 to prove Theorem 1.5 for the weights $\gamma^{(a,b)}$ and $\delta^{(a',b')}$, and give further results on the left- and right-eigenvectors of the transition matrix $P(\gamma^{(a,b)})$. We end this section with the proof of Corollary 1.10. In §6 we prove Corollary 1.6 and in §7 we prove Theorem 1.8. In §8 we give a Hilbert space setting for involutive walks on the interval [0,1] and in §9 we prove analogues of the main results for the $\gamma^{(a,b)}$ -weighted discrete walks. We end by comparing the spectral behaviour in the discrete and continuous cases.

2. The global anti-diagonal eigenvalue property

In this section we prove Theorem 2.9 characterizing the binomial transforms H^{λ} in (1.6) in terms of Definition 2.7. We also give a self-contained proof of the connection between the binomial transform of a matrix and the global anti-diagonal eigenvalue property, first proved in [6], and prove some simpler results sufficient for the applications to probability below.

2.1. **Preliminaries.** Recall that B(n) is the $n \times n$ Pascal's Triangle matrix defined by $B(n)_{xy} = \binom{x}{y}$ and that J(n) is the $n \times n$ matrix defined in Iverson

bracket notation by $J(n)_{xy} = [x = y^*]$. For $e \in \{0, 1, ..., n\}$, let V_e be the e-dimensional subspace of \mathbb{R}^n of all row vectors (f(0), ..., f(n-1)) where f is a polynomial of degree strictly less than e. Let

(2.1)
$$v(d) = \left(\begin{pmatrix} 0 \\ d \end{pmatrix}, \begin{pmatrix} 1 \\ d \end{pmatrix}, \dots, \begin{pmatrix} n-1 \\ d \end{pmatrix} \right)^t.$$

Thus B(n) has columns $v(0), \ldots, v(n-1)$.

Lemma 2.1.

- (i) $B(n)_{zw}^{-1} = (-1)^{z+w} {z \choose w};$
- (ii) $V_e = \langle v(0), \dots, v(e-1) \rangle$ for each $e \in \{0, \dots, n\}$;
- (iii) $J(n)v(d) \in (-1)^d v(d) + \langle v(0), \dots, v(d-1) \rangle$ for each $d \in \{0, \dots, n-1\}$;

(iv)
$$\left(B(n)^{-1}J(n)B(n)\right)_{xy} = (-1)^x {x^* \choose y^*}.$$

Proof. For (i) we calculate

$$\sum_{w=0}^{n-1} {x \choose w} (-1)^{x+w} B(n)_{wy} = (-1)^x \sum_{w=0}^{n-1} (-1)^w {x \choose w} {w \choose y}$$
$$= (-1)^x {x \choose y} \sum_{w=0}^{n-1} (-1)^w {x - y \choose w - y}$$
$$= [x = y]$$

where in the final step we use the corollary of the Binomial Theorem that $\sum_{z=0}^{m} (-1)^z \binom{m}{z} = [m=0]$. Since the binomial coefficient $\binom{X}{d}$ is a polynomial with leading term $X^d/d!$, we have (ii). For (iii) we note that J(n)v(d) has as its entries $\binom{n-1-X}{d}$ evaluated at $0,\ldots,n-1$. Since this binomial coefficient is a polynomial of degree d with leading term $(-1)^d X^d/d!$, (iii) follows from (ii). Finally for (iv) we apply (i), the identity $\binom{a}{b} = \binom{-b-1}{a-b}(-1)^{a-b}$ valid for $a \ge b$, Vandermonde's convolution, and finally $\binom{b-a-1}{b}(-1)^b = \binom{a}{b}$ again valid for $a \ge b$, to get

$$(B(n)^{-1}J(n)B(n))_{xy} = \sum_{z} (-1)^{x+z} \binom{x}{z} \binom{n-1-z}{y}$$

$$= \sum_{z} (-1)^{n-1+x+y} \binom{x}{z} \binom{-y-1}{n-1-z-y}$$

$$= (-1)^{n-1+x+y} \binom{x-y-1}{n-1-y}$$

$$= (-1)^x \binom{n-1-x}{n-1-y}$$

as required.

2.2. A sufficient condition for the global anti-diagonal eigenvalue **property.** We use the following lemma and proposition in §4 and §5 to show that the down-step matrices of the involutive walks for the binomial weights $\gamma^{(a,b)}$, $\gamma^{(c)}$ and $\delta^{(a',b')}$ are binomial transforms, and so have the global anti-diagonal eigenvalue property.

Lemma 2.2. Let L be a real lower-triangular matrix with diagonal entries $\lambda_0, \ldots, \lambda_{n-1}$ read from top to bottom. The following are equivalent

- (i) L is equal to the binomial transform H^{λ} ;
- (ii) $L = B(n)\operatorname{Diag}(\lambda_0, \dots, \lambda_{n-1})B(n)^{-1};$
- (iii) $Lv(d) = \lambda_d v(d)$ for each $d \in \{0, \dots, n-1\}$.
- (iv) $L_{xy} = \binom{x}{y} \sum_{e=0}^{x-y} (-1)^e \binom{x-y}{e} \lambda_{y+e} \text{ for } 0 \leq y \leq x < n.$

Proof. Conditions (i) and (ii) are immediately equivalent by the definition of H^{λ} . Since B(n) has columns $v(0), v(1), \ldots, v(n-1)$, (iii) holds if and only if $LB(n) = B(n) \text{Diag}(\lambda_0, \ldots, \lambda_{n-1})$, hence if and only if (ii) holds. Finally, by Lemma 2.1(i), the entry H^{λ}_{xy} is

$$\sum_{w=0}^{n-1} {x \choose w} \lambda_w (-1)^{w+y} {w \choose y} = \sum_{w=y}^{n-1} {x \choose y} (-1)^{w+y} {x-y \choose w-y} \lambda_w$$
$$= {x \choose y} \sum_{e=0}^{x-y} (-1)^e {x-y \choose e} \lambda_{y+e}$$

as required for (iv).

Proposition 2.3. Suppose that any of the equivalent conditions in Lemma 2.2 holds for the $n \times n$ lower-triangular matrix L. Then

- (i) $LJ(n)v(d) \in (-1)^d \lambda_d v(d) + V_{d-1};$
- (ii) LJ(n) preserves each subspace V_d and acts as $(-1)^d \lambda_d$ on V_{d+1}/V_d ;
- (iii) The matrix representing LJ(n) in the basis $v(0), v(1), \ldots, v(n-1)$ is upper-triangular with diagonal entries $\lambda_0, -\lambda_1, \ldots, (-1)^{n-1}\lambda_{n-1}$ read from top to bottom.
 - (iv) L has the global anti-diagonal eigenvalue property.

Proof. By Lemma 2.2(iii), $Lv(d) = \lambda_d v(d)$. By Lemma 2.1(ii) and (iii), $v(d) \in V_{d+1}$ and $J(n)v(d) = (-1)^d v(d) + w$ where $w \in V_d$. Hence $LJv(d) \in (-1)^d \lambda_d v(d) + V_d$, as required for (i). Since $v(d) + V_d$ spans the quotient space V_{d+1}/V_d , (ii) follows at once; (iii) simply restates (ii) in matrix language. From (iii) we see that LJ(n) has eigenvalues $\lambda_0, -\lambda_1, \ldots, (-1)^{n-1}\lambda_{n-1}$ read from top-to-bottom. Hence L has the anti-diagonal eigenvalue property. By applying the canonical projection $\mathbb{R}^n \to \mathbb{R}^m$ to each subspace V_d , we see that all these results hold for each $m \times m$ top-left submatrix of L. Hence L has the global anti-diagonal eigenvalue property.

2.3. A generic characterisation of the global anti-diagonal eigenvalue property. We now give a self-contained proof of Theorem 2.17 in [6]. The matrix C in our proof plays a similar role to the matrix in Lemma 2.13 in [6], but is more easily visualised. Stated so as to make clear that the exceptional set is an algebraic variety, Theorem 2.17 in [6] is as follows.

Theorem 2.4 (Ochiai, Sasada, Shirai and Tsuboi). There is a homogeneous polynomial $f(x_0, \ldots, x_{n-2})$ of degree (n-1)(n-2)/2 such that, provided $f(\lambda_0, \ldots, \lambda_{n-2}) \neq 0$, a lower-triangular matrix L with eigenvalues $\lambda_0, \ldots, \lambda_{n-2}, \lambda_{n-1}$ read from top to bottom has the global anti-diagonal eigenvalue property if and only if $L = H^{\lambda}$.

Proof. The 'if' direction holds by Proposition 2.3(iv). Suppose that L is a real lower-triangular $n \times n$ matrix with eigenvalues $\lambda_0, \ldots, \lambda_{n-1}$ having the global anti-diagonal eigenvalue property. Working by induction on n we may assume that the matrix obtained from L by deleting its bottom row and rightmost column is $H^{(\lambda_0,\ldots,\lambda_{n-2})}$. Since $H^{(\lambda_0,\ldots,\lambda_{n-2})}$ is obtained by the same deletions from $H^{(\lambda_0,\ldots,\lambda_{n-2},\lambda_{n-1})}$, we may write $L=H^{\lambda}+E$ where the $n \times n$ matrix E is zero except in its bottom row. Let the bottom row of E be $\varepsilon_0,\ldots,\varepsilon_{n-1}$ read left to right. It suffices to show that $\varepsilon_z=0$ for each z.

By the assumption that L has eigenvalues $\lambda_0, \ldots, \lambda_{n-1}$ we have $\operatorname{tr} L = \lambda_0 + \lambda_1 + \cdots + \lambda_{n-1}$. Since $\operatorname{tr} L = \operatorname{tr}(H^{\lambda} + E) = \lambda_0 + \lambda_1 + \cdots + \lambda_{n-1} + \varepsilon_{n-1}$, it follows that $\varepsilon_{n-1} = 0$. Similarly, since the matrices LJ(n) and $H^{\lambda}J(n)$ have eigenvalues $\lambda_0, -\lambda_1, \ldots, (-1)^{n-1}\lambda_{n-1}$, we have $\operatorname{tr} LJ(n) = \operatorname{tr} H^{\lambda}J(n)$; now since $LJ(n) = H^{\lambda}J(n) + EJ(n)$ and EJ(n) has bottom row entries $\varepsilon_{n-1}, \ldots, \varepsilon_0$, it follows that $\varepsilon_0 = 0$.

Consider the matrix LJ(n) - XI, where X is an indeterminate. Since L has the anti-diagonal eigenvalue property, $\det(LJ(n) - XI) = \det(H^{\lambda}J(n) - XI)$. By the multilinearity of the determinant,

$$\det(LJ(n) - XI) - \det(H^{\lambda}J(n) - XI) = \det N$$

where N is the matrix obtained from LJ(n)-XI by subtracting the bottom row of $H^{\lambda}J(n)$; this leaves the bottom row of EJ(n). For instance if n=5 then N is as shown below.

$$\begin{pmatrix} -X & \cdot & \cdot & \cdot & \lambda_0 \\ \cdot & -X & \cdot & \lambda_1 & \lambda_0 - \lambda_1 \\ \cdot & \cdot & \lambda_2 - X & 2(\lambda_1 - \lambda_2) & \lambda_0 - 2\lambda_1 + \lambda_2 \\ \cdot & \lambda_3 & 3(\lambda_2 - \lambda_3) & 3(\lambda_0 - 2\lambda_1 + \lambda_2) - X & \lambda_0 - 3\lambda_1 + 3\lambda_2 - \lambda_3 \\ 0 & \varepsilon_3 & \varepsilon_2 & \varepsilon_1 & 0 \end{pmatrix}$$

(The entries in N can be computed using Lemma 2.2(iv), or, more conveniently for hand calculations, by Lemma 6.1 below.)

For each $d \in \{0, 1, \ldots, n-1\}$, the coefficient of X^d in $\det N$ is a linear combination $f_1\varepsilon_1+\cdots+f_{n-2}\varepsilon_{n-2}$ where each f_y is a homogeneous polynomial in $\lambda_0, \ldots, \lambda_{n-2}$ of degree n-1-d. Since the coefficients of X^0 and X^{n-1} are zero, we need only consider $d \in \{1, \ldots, n-2\}$. Let C be the $(n-2) \times (n-2)$ matrix with rows and columns labelled by $\{1, \ldots, n-2\}$ such that C_{dy} is the coefficient of $\varepsilon_y X^d$ in $\det N$. There is a non-zero solution $(\varepsilon_1, \ldots, \varepsilon_{n-2})$ to $\det N = 0$ if and only if $\det C = 0$. We can therefore take

the required multivariable polynomial f to be $\det C$, provided $\det C$ is not the zero polynomial in the eigenvalues $\lambda_0, \ldots, \lambda_{n-1}$.

This is surprisingly non-trivial to prove. It suffices to show that $\det C$ is non-zero for one specialization of the eigenvalues. We take $\lambda_x = x$ for each x. (The choice for λ_{n-1} is irrelevant.) After this specialization, the matrices N and C are shown left and right below when n=7.

$$\begin{pmatrix} -X & \cdot & \cdot & \cdot & \cdot & \cdot & 0 \\ \cdot & -X & \cdot & \cdot & \cdot & 1 & -1 \\ \cdot & \cdot & -X & \cdot & 2 & -2 & 0 \\ \cdot & \cdot & \cdot & 3 - X & -3 & 0 & 0 \\ \cdot & \cdot & 4 & -4 & -X & 0 & 0 \\ \cdot & 5 & -5 & 0 & 0 & -X & 0 \\ 0 & \varepsilon_5 & \varepsilon_4 & \varepsilon_3 & \varepsilon_2 & \varepsilon_1 & -X \end{pmatrix} \quad \begin{matrix} X \\ X^2 \\ X^3 \\ X^4 \\ X^5 \\ X^5 \\ X^6 \\ X^5 \\ X^6 \\ X^5 \\ X^6 \\ X^6 \\ X^5 \\ X^6 \\ X$$

Observe that, after a column permutation C is upper-triangular, with non-zero diagonal entries, and so has non-zero determinant.

In general, to obtain a non-zero contribution to the coefficient of ε_y by expanding det N we must take ε_y from the bottom row, -X from the top row, and -1 from row 1. Thus C_{dy} is the coefficient of $\varepsilon_y X^{d-1}$ in the minor obtained by deleting the first two rows and the first and last columns from N. We may rearrange the rows and the rows and columns of this minor matrix so that (using the original numbering of $\{0,\ldots,n-1\}$ for both), the rows come in order $n-2,2,n-3,3,\ldots,\lfloor\frac{n}{2}\rfloor,n-1$ and the columns in order $1,n-2,2,n-3,\ldots,\lfloor\frac{n}{2}\rfloor$. This gives an $(n-2)\times(n-2)$ matrix N' that is upper-triangular and tridiagonal, except in its bottom row which has entries $\varepsilon_{n-2},\varepsilon_1,\varepsilon_{n-3},\varepsilon_2,\ldots,\varepsilon_{\lfloor\frac{n-1}{2}\rfloor}$. (To see this, note that the columns with highest entry in the same row are now adjacent, as are the rows with left-most entry in the same column.) By this choice of order, the entries of N involving X appear on the super-diagonal of N'. More precisely, labelling rows and columns from 0 as usual, if x < n-4 is even then

$$N'_{xy} = \begin{cases} n - 2 - \frac{x}{2} & \text{if } y = x \\ -X & \text{if } y = x + 1 \\ -(n - 2 - \frac{x}{2}) & \text{if } y = x + 2 \end{cases}$$

and if x < n - 4 is odd then

$$N'_{xy} = \begin{cases} -(1 + \frac{x+1}{2}) & \text{if } y = x \\ -X & \text{if } y = x+1 \\ 1 + \frac{x+1}{2} & \text{if } y = x+2 \end{cases}$$

and all other entries of N' in these rows are 0. Row n-4 of N', just above the bottom row, ends with the two non-zero entries $\pm \lfloor \frac{n}{2} \rfloor$, $-X \mp \lfloor \frac{n}{2} \rfloor$, the signs depending on the parity of n. For example, if n=7 then N' is as

shown below.

$$\begin{pmatrix}
5 & -X & -5 & 0 & 0 \\
0 & -2 & -X & 2 & 0 \\
0 & 0 & 4 & -X & -4 \\
0 & 0 & 0 & -3 & 3 - X \\
\varepsilon_5 & \varepsilon_1 & \varepsilon_4 & \varepsilon_2 & \varepsilon_3
\end{pmatrix}$$

The minor matrix obtained by deleting the bottom row and final column from N' is upper-triangular; its determinant is the coefficient of $\varepsilon_{\lfloor \frac{n-1}{2} \rfloor}$ and has degree 0 in X. Generally, if ε_y comes r-th from the end in the order ε_{n-2} , $\varepsilon_1, \, \varepsilon_{n-3}, \, \varepsilon_2, \, \ldots, \, \varepsilon_{\lfloor \frac{n-1}{2} \rfloor}$, then the corresponding minor matrix of N' has an $(r-1) \times (r-1)$ bottom right matrix with determinant of degree r-1 in X and the minor is the product of this by a non-zero constant. Therefore the polynomial coefficients of $\varepsilon_{n-2}, \, \varepsilon_1, \, \varepsilon_{n-3}, \, \varepsilon_2, \, \ldots, \, \varepsilon_{\lfloor \frac{n-1}{2} \rfloor}$ are of degrees $n-3, \, n-4, \, \ldots, \, 1, \, 0$, respectively. It follows that the matrix C can be rearranged (by the reverse of this ordering of the ε_y) so that it is upper-triangular with non-zero diagonal entries. Hence $\det C \neq 0$.

The following example extends those in [6] to show that the exceptional set in Theorem 2.4 is necessary even if we add the condition that the lower-triangular matrix has distinct eigenvalues, or is stochastic.

Example 2.5. It is a routine calculation to show that the lower-triangular matrices L and H left and middle below both have the anti-diagonal eigenvalue property; an explicit basis of right-eigenvectors of HJ(5) is given in the far right matrix. By Lemma 2.2(iv), the 3×3 top-left submatrix of L is $H^{(1,\frac{2}{3},\frac{1}{4})}$, and similarly, the 4×4 top-left submatrix of H is $H^{(1,\frac{1}{2},\frac{1}{2}-\tau,\frac{1}{2}-\tau)}$. Therefore L and H have the global anti-diagonal eigenvalue property.

$$\begin{pmatrix} 1 & \cdot & \cdot & \cdot \\ \frac{1}{3} & \frac{2}{3} & \cdot & \cdot \\ -\frac{1}{12} & \frac{5}{6} & \frac{1}{4} & \cdot \\ -\frac{9}{20} & \frac{11}{10} + \frac{4}{5}\tau & \frac{3}{20} + \frac{1}{5}\tau & \frac{1}{5} \end{pmatrix} \quad \begin{pmatrix} 1 & \cdot & \cdot & \cdot & \cdot \\ \frac{1}{2} & \frac{1}{2} & \cdot & \cdot & \cdot & \cdot \\ \frac{1}{2} - \tau & 2\tau & \frac{1}{2} - \tau & \cdot & \cdot & \cdot \\ \frac{1}{2} - 2\tau & 3\tau & 0 & \frac{1}{2} - \tau & \cdot \\ \frac{1}{2} - 4\tau & 5\tau & 0 & \tau & \frac{1}{2} - 2\tau \end{pmatrix} \quad \begin{pmatrix} 1 & -2 & 0 & -4 & -4\tau(10\tau - 3) \\ 1 & -\frac{5}{4} & 0 & -4 & -28\tau^2 + 11\tau - 1 \\ 1 & -\frac{1}{2} & 1 & -2\tau - 1 & (4\tau - 1)\left(4\tau^2 - 6\tau + 1\right) \\ 1 & \frac{1}{4} & 0 & 2 - 4\tau & (4\tau - 1)\left(8\tau^2 - 5\tau + 1\right) \\ 1 & 1 & 0 & 2 - 4\tau & 2\tau(4\tau - 1)(10\tau - 3) \end{pmatrix}$$

Again by Lemma 2.2(iv), if $\tau \neq 0$ then the bottom row of L does not agree with $H^{(1,\frac{2}{3},\frac{1}{4},\frac{1}{5})}$, and similarly the bottom row of H does not agree with $H^{(1,\frac{1}{2},\frac{1}{2}-\tau,\frac{1}{2}-\tau,\frac{1}{2}-2\tau)}$. Therefore these matrices have the global anti-diagonal eigenvalue property but are not of the form H^{λ} .

The authors know of no example satisfying both conditions simultaneously. An explicit calculation of the determinant $\det C$ in the previous proof can be used to prove the conjecture below for $n \leq 4$.

Conjecture 2.6. A stochastic matrix with distinct eigenvalues has the global anti-diagonal eigenvalue property if and only if it is of the form H^{λ} .

We offer this as an open problem, related to Conjecture 1.9.

2.4. Anti-diagonal conjugators and a characterisation of the matrices H^{λ} . Let L be a diagonalizable lower-triangular matrix having diagonal entries $\lambda_0, \ldots, \lambda_{n-1}$. There is a basis $u(0), \ldots, u(n-1)$ of eigenvectors of L such that the matrix Q with columns $u(0), \ldots, u(n-1)$ is lower-triangular. The matrix representing LJ(n) in this basis is

$$Q^{-1}LJ(n)Q = (Q^{-1}LQ)(Q^{-1}J(n)Q) = \text{Diag}(\lambda_0, \dots, \lambda_{n-1})Q^{-1}J(n)Q.$$

Thus if $Q^{-1}J(n)Q$ is upper-triangular with alternating sign entries, then L has the anti-diagonal eigenvalue property. Motivated by this observation we make the following definition, working over an arbitrary infinite field \mathbb{F} .

Definition 2.7. A lower-triangular $n \times n$ matrix Q with entries in \mathbb{F} is an anti-diagonal conjugator if $Q^{-1}J(n)Q$ is upper-triangular with diagonal entries $(Q^{-1}J(n)Q)_{xx} = (-1)^x$. It is a global anti-diagonal conjugator if the same holds for all the top-left $m \times m$ submatrices of Q, replacing J(n) with J(m). A lower-triangular matrix L has global anti-triangular eigenbasis action if it has the columns of a global anti-diagonal conjugator as its eigenvectors.

By Lemma 2.2(iv), $B(n)^{-1}J(n)B(n)$ is upper-triangular for each n. Therefore B(n) is a global anti-diagonal conjugator, and the binomial transform $H^{\lambda} = B(n)\operatorname{Diag}(\lambda_0, \ldots, \lambda_{n-1})B(n)^{-1}$ has a global anti-diagonal eigenbasis. By the following proposition, up to column scaling of B(n), this is the only family of examples. The final remark in §9.3 gives an intuitive interpretation of this proposition.

Proposition 2.8. A lower-unitriangular $n \times n$ matrix Q is a global anti-diagonal conjugator if and only if Q = B(n).

Proof. We work by induction on n. The result is obvious if n=1. For the inductive step, we suppose that Q is an $(n+1)\times(n+1)$ global anti-diagonal conjugator. The $n\times n$ top-left submatrix of Q is a global anti-diagonal conjugator, and so by induction, it is B(n). Let $Q_{nd}=\alpha_d$ for $d\leqslant n$. Let $u(d)\in\mathbb{F}^{n+1}$ be column d of Q of length n+1 and let $v(d)\in\mathbb{F}^n$ be column d of B(n) of length n. For $d\in\{0,\ldots,n-1\}$ we have $u(d)_y=v(d)_y$ if y< n and $u(d)_n=\alpha_d$. By the hypothesis that J(n+1)Q=QU where U is upper-triangular with the specified diagonal entries, we have

$$(2.2) J(n+1)u(d) \in (-1)^d u(d) + \langle u(0), \dots, u(d-1) \rangle.$$

We shall deduce from (2.2) that $\alpha_d = \binom{n}{d}$. Since J(n+1)u(0) = u(0) we have $\alpha_0 = u(0)_n = u(0)_0 = 1$. Moreover, $\alpha_n = Q_{nn} = 1$ since Q_{nn} is uni-triangular. Let $1 \leq d < n$ and let μ_0, \ldots, μ_{d-1} be the unique scalars given by (2.2) such that

(2.3)
$$J(n+1)u(d) = (-1)^d u(d) + \sum_{x=0}^{d-1} \mu_x u(x).$$

Projecting down to \mathbb{F}^n by taking the entries in positions $0, 1, \ldots, n-1$ on either side we may replace u(x) with v(x) for each $x \leq d$ to get

(2.4)
$$\left(\alpha_d, \binom{n-1}{d}, \dots, \binom{1}{d}\right)^t = (-1)^d v(d) + \sum_{x=0}^{d-1} \mu_x v(x).$$

By Lemma 2.1(iv), the matrix representing J(n) in the basis $v(0), \dots v(n-1)$ is upper-triangular with entries $(-1)^x \binom{n-1-x}{n-1-y}$. Hence

(2.5)
$$J(n)v(d) = (-1)^{d}v(d) + \sum_{x=0}^{d-1} (-1)^{x} \binom{n-1-x}{n-1-d} v(x).$$

This is a somewhat similar relation to (2.4). Indeed $J(n)v(d)_x = \binom{n-1-x}{d}$, and adding $\binom{n-1-x}{d-1}$ gives $\binom{n-x}{d}$, agreeing with the left-hand side of (2.4) for $x \ge 1$. Observe that $\binom{n-1-x}{d-1} = \left(J(n)v(d-1)\right)_x$ and, by replacing d with d-1 in the previous displayed equation,

(2.6)
$$J(n)v(d-1) = (-1)^{d-1}v(d-1) + \sum_{x=0}^{d-2} (-1)^x \binom{n-1-x}{n-d} v(x).$$

Adding (2.5) and (2.6) and using that $\binom{n-1-x}{d} + \binom{n-1-x}{d-1} = \binom{n-x}{d}$ on the left-hand side and $\binom{n-1-x}{n-1-d} + \binom{n-1-x}{n-d} = \binom{n-x}{n-d}$ on the right-hand side gives

$$\left(\binom{n}{d}, \binom{n-1}{d}, \dots, \binom{1}{d}\right)^{t} = J(n)v(d) + J(n)v(d-1)
= (-1)^{d}v(d) + \sum_{x=0}^{d-1} (-1)^{x} \binom{n-x}{n-d} v(x).$$

Taking the difference of (2.4) and (2.7) we get

$$\left(\alpha_d - \binom{n}{d}, 0, \dots, 0\right)^{t} = \sum_{x=0}^{d-1} \left(\mu_x - (-1)^x \binom{n-x}{n-d}\right) v(x).$$

The function f defined by f(0) = 1 and f(x) = 0 for $x \in \{1, ..., n-1\}$ has n-1 zeros and is therefore a polynomial of degree n-1. Thus, if non-zero, the left-hand side is in V_n but not in V_e for any e < n. On the other hand, the right-hand side is clearly in V_d . Since d < n, we conclude that $\alpha_d = \binom{n}{d}$, as required.

Theorem 2.9. Let L be a lower-triangular $n \times n$ matrix with entries in \mathbb{F} having diagonal entries $\lambda_0, \ldots, \lambda_{n-1}$ read from top-bottom. Then L has global anti-diagonal eigenbasis action if and only if L is equal to the binomial transform $H^{(\lambda_0,\ldots,\lambda_{n-1})}$.

Proof. This is immediate from Definition 2.7 and Proposition 2.8

3. Factorizable weights and proofs of Theorems 1.3 and 1.4

The second hypothesis in Theorem 1.3, that $\gamma_{[x,x]} > 0$ for all $x \in n$ and $\gamma_{[x-1,x]} > 0$ for all non-zero $x \in n$, implies that the steps $x \mapsto x^*$ and $x \mapsto (x-1)^*$ have positive probability in the γ -weighted involutive walk for all relevant x. For later use in §7, we take a weaker hypothesis below.

Lemma 3.1. In an involutive walk on n, if x^* is accessible from x for all $x \in n$ and $(x-1)^*$ is accessible from x for all non-zero $x \in n$ then the walk is irreducible.

Proof. The hypothesis implies that, if the involutive walk is allowed to take multiple steps, then each transition in

$$0 \rightsquigarrow n-1 \rightsquigarrow 1 \rightsquigarrow n-2 \rightsquigarrow 2 \rightsquigarrow \cdots \rightsquigarrow n-2 \rightsquigarrow 1 \rightsquigarrow n-1 \rightsquigarrow 0$$

has non-zero probability. Therefore all states communicate.

We are now ready to prove Theorem 1.3.

Proof of Theorem 1.3. If (i) holds then, by (1.1), the steps $x \mapsto 0^* = n-1$ and $n-1 \mapsto x^*$ have non-zero probability for all $x \in n$. Hence the involutive walk is irreducible. Moreover, since $n-1 \mapsto 0^* = n-1$ is a possible step, it follows from a standard criterion that the involutive walk is aperiodic. If (ii) holds then, again by (1.1), the steps $x \mapsto x^*$ and $x \mapsto (x-1)^*$ have non-zero probabilities and so, by Lemma 3.1, the involutive walk is again irreducible. Moreover, if n=2m then $m \mapsto (m-1)^* = m$ is a possible step and if n=2m+1 then $m \mapsto m^* = m$ is a possible step. Hence the involutive walk is aperiodic. Thus in either case there is a unique invariant distribution to which the involutive walk converges by the ergodic theorem.

We now turn to reversibility. The detailed balance equations for an invariant probability distribution π on n are $\pi_x P(\gamma)_{xz} = \pi_z P(\gamma)_{zx}$. Equivalently,

(3.1)
$$\pi_x \frac{\gamma_{[z^*,x]}}{N(\gamma)_x} = \pi_z \frac{\gamma_{[x^*,z]}}{N(\gamma)_z}$$

for $x, z \in n$. Since $\gamma_{\varnothing} = 0$ and $[z^{\star}, x]$ is non-empty if and only if $[x^{\star}, z]$ is non-empty, we may assume that $z^{\star} \leq x$ in this equation.

Lemma 3.2.

- (i) If α is an atomic weight then the α -weighted involutive walk is reversible with invariant distribution $\pi(\alpha)$ such that $\pi(\alpha)_x \propto \alpha_{x^*} N(\alpha)_x$.
- (ii) If β is a \star -symmetric weight then the β -weighted involutive walk is reversible with an invariant distribution $\pi(\beta)$ such that $\pi(\beta)_x \propto N(\beta)_x$.

Proof. When α is atomic the equations (3.1) simplify to $\pi_x \alpha_{z^*}/N(\alpha)_x = \pi_z \alpha_{x^*}/N(\alpha)_z$ for $x, z \in n$ such that $z^* \leq x$. Clearly one solution has each π_x proportional to $\alpha_{x^*}N(\alpha)_x$. If β is \star -symmetric then, by definition, $\beta_{[z^*,x]} = \beta_{[x^*,z]}$, and so one solution to (3.1) has each π_x proportional to $N(\beta)_x$. \square

Since atomic weights are strictly positive, both invariant distributions are strictly positive. Reversibility is preserved by products of weights.

Lemma 3.3. Let α and β be weights. If the α -weighted and β -weighted involutive walks on n are reversible with respect to invariant distributions proportional to ϑ and ϕ , respectively, then the $\alpha\beta$ -weighted involutive walk on n is reversible with respect to an invariant distribution proportional to $\vartheta_x \phi_x N(\alpha\beta)_x / N(\alpha)_x N(\beta)_x$.

Proof. Multiplying the two cases of (3.1) for α and β we obtain

$$\vartheta_x \phi_x \frac{\alpha_{[z^*,x]} \beta_{[z^*,x]}}{N(\alpha)_x N(\beta)_x} = \vartheta_z \phi_z \frac{\alpha_{[x^*,z]} \beta_{[x^*,z]}}{N(\alpha)_z N(\beta)_z}$$

for all $x, z \in n$. Therefore $\pi_x = \vartheta_x \phi_x N(\alpha \beta)_x / N(\alpha)_x N(\beta)_x$ solves (3.1) for the weight $\alpha \beta$.

We are now ready to prove Theorem 1.4.

Proof of Theorem 1.4. Suppose that $\gamma = \alpha \beta$ factorizes as in the theorem. By Lemma 3.2, the α - and β -weighted involutive walks are reversible with respect to invariant distributions proportional to $\alpha_{x^*}N(\alpha)_x$ and $N(\beta)_x$, respectively. Therefore, by Lemma 3.3, the γ -weighted involutive walk is reversible with respect to an invariant distribution proportional to

$$\alpha_{x^*} N(\alpha)_x N(\beta)_x \times \frac{N(\alpha\beta)_x}{N(\alpha)_x N(\beta)_x} = \alpha_{x^*} N(\alpha\beta)_x.$$

By Theorem 1.3, this invariant distribution is unique.

Suppose conversely that the γ -weighted involutive walk is reversible with respect to an invariant distribution π . By the hypothesis of Theorem 1.4 and Theorem 1.3, the chain is irreducible and ergodic, and so this invariant distribution is strictly positive. We must define an atomic weight α so that β , defined by $\beta = \gamma/\alpha$, is \star -symmetric. Taking $y = z^{\star}$ in (3.1) and rearranging we obtain

(3.2)
$$\frac{N(\gamma)_{y^{\star}}}{\pi_{y^{\star}}}\gamma_{[y,x]} = \frac{N(\gamma)_{x^{\star}}}{\pi_{x^{\star}}}\gamma_{[x^{\star},y^{\star}]}$$

for all $x, y \in n$. Noting that the fractions on each side depend only on y^* and only on x^* , respectively, we are led to define an atomic weight α by $\alpha_{[y,x]} = \frac{\pi_{y^*}}{N(\gamma)_{y^*}}$. Now by (3.2),

$$\beta_{[y,x]} = \frac{\gamma_{[y,x]}}{\alpha_y} = \frac{N(\gamma)_{y^{\star}}}{\pi_{y^{\star}}} \gamma_{[y,x]} = \frac{N(\gamma)_{x^{\star}}}{\pi_{x^{\star}}} \gamma_{[x^{\star},y^{\star}]} = \frac{\gamma_{[x^{\star},y^{\star}]}}{\alpha_{y^{\star}}} = \beta_{[x^{\star},y^{\star}]}$$

so β is \star -symmetric, as required.

4. The $\gamma^{(c)}$ -weighted involutive walk

Let $c \in \mathbb{R}^{>0}$. Recall that $\gamma^{(c)}$ is the weight defined by $\gamma^{(c)}_{[y,x]} = {x \choose y} c^{x-y}$ for $y \leq x$. Since $N(\gamma^{(c)})_x = (c+1)^x$ we have

(4.1)
$$H(\gamma^{(c)})_{xy} = {x \choose y} \frac{c^{x-y}}{(c+1)^x}.$$

In particular, $H(\gamma^{(1)})$ is the Pascal's Triangle matrix with rows scaled so that it is stochastic. Observe that the eigenvalues of $H(\gamma^{(c)})$, and so the anti-diagonal entries of $P(\gamma^{(c)})$, are $1/(c+1)^d$ for $d \in \{0, \ldots, n-1\}$.

4.1. Stochastic properties of $P(\gamma^{(c)})$. Since $\gamma^{(c)}$ is strictly positive, Theorem 1.3 implies that the $\gamma^{(c)}$ -weighted involutive walk is irreducible, recurrent and ergodic with a unique invariant distribution. To find the invariant distribution we first show that $\gamma^{(c)}$ is factorizable. (This is motivated by the interpretation of $\gamma^{(c)}$ as a limiting case of the clearly factorizable weights $\gamma^{(a,ac)}$.) From the required form $\binom{x}{y}c^{x-y} = \alpha_y\beta_{[y,x]}$, where β is \star -symmetric, we must define α_y so that $\alpha_yx!c^{x-y}/y!(x-y)!$ is invariant under $(x,y)\mapsto (y^\star,x^\star)$. This already holds for c^{x-y} and (x-y)!, and similarly to the proof of Theorem 1.4, we may cancel y! and introduce the necessary factor of $y^\star!$ by taking α_y proportional to 1/y!(n-1-y)!. We therefore define $\alpha_y=\binom{n-1}{y}$ as a weight with domain n, and find that

(4.2)
$$\beta_{[y,x]} = \frac{\binom{x}{y}c^{x-y}}{\binom{n-1}{y}} = \frac{x!y^{*}!}{(x-y)!(n-1)!}c^{x-y}$$

is indeed \star -symmetric. Theorem 1.4 now implies that the $\gamma^{(c)}$ -weighted involutive walk on n is reversible with unique invariant distribution proportional to $\alpha_{x^{\star}}N(\gamma^{(c)})_x$; that is

$$\pi_x \propto \binom{n-1}{x} \left(\frac{c+1}{c}\right)^x$$
.

Summing the right-hand side using the binomial theorem we get the explicit formula

$$\pi_x = \binom{n-1}{x} \frac{(c+1)^x c^{n-1-x}}{(2c+1)^{n-1}}.$$

4.2. **Spectrum of** $\gamma^{(c)}$. We now prove that the eigenvalues of $P(\gamma^{(c)})$ are $(-1)^d/c^d$ for $d \in \{0, \ldots, n-1\}$ using the condition in Lemma 2.2(iii) for $H(\gamma^{(c)})$ to be a binomial transform. Recall from (2.1) that $v(d) \in \mathbb{R}^n$ is the column vector with entries $v(d)_x = {x \choose d}$.

Lemma 4.1. Let $c \in \mathbb{R}^{>0}$. Then

$$H(\gamma^{(c)})v(d) = v(d)/(c+1)^d$$
.

Proof. Since $H(\gamma^{(c)})_{xy} = \binom{x}{y} \frac{c^{x-y}}{(c+1)^x}$ we have

$$\begin{split} \left(H(\gamma^{(c)})v(d)\right)_x &= \sum_{y=0}^{n-1} \binom{x}{y} \frac{c^{x-y}}{(c+1)^x} \binom{y}{d} = \binom{x}{d} \frac{c^{x-d}}{(c+1)^x} \sum_{y=0}^{n-1} \binom{x-d}{y-d} \frac{1}{c^{y-d}} \\ &= \binom{x}{d} \frac{c^{x-d}}{(c+1)^x} \Big(1 + \frac{1}{c}\Big)^{x-d} = \binom{x}{d} \frac{1}{(c+1)^d} = \frac{v(d)_x}{(c+1)^d} \end{split}$$
 as required.

Hence $H(\gamma^{(c)})$ is the binomial transform $H^{\lambda^{(c)}}$, where $\lambda^{(c)}_d=1/(c+1)^d$ for each $d \in \{0, \dots, n-1\}$, and, by Proposition 2.3(iii), $P(\gamma^{(c)})$ has eigenvalues $(-1)^d/(c+1)^d$ for $d \in \{0,\ldots,n-1\}$, as required. This completes the proof of Theorem 1.5 for $\gamma^{(c)}$.

5. The $\gamma^{(a,b)}$ -weighted and $\delta^{(a',b')}$ -weighted involutive walks

We now adopt a similar strategy to the previous section to prove Theorem 1.5 for the $\gamma^{(a,b)}$ and $\delta^{(a',b')}$ involutive walks. It is convenient in this section to extend this definition of $\gamma^{(a,b)}$ to define $\gamma^{(a,b)}_{[y,x]} = \binom{a+y}{y} \binom{b+x-y}{x-y}$ for all $a, b \in \mathbb{R}$ and all $x, y \in \mathbb{N}$ with $y \leq x$.

5.1. **Preliminaries on multisubsets.** A combinatorial interpretation of $\gamma^{(a,b)}$ is helpful and motivates many steps below. Recall that $\binom{m+c-1}{c}$, or equally $(-1)^c {-m \choose c}$, is the number of c-multisubsets of a set of size m, or equivalently, the number of chains $0 \le y_1 \le \ldots \le y_c \le m-1$ in the ordinal m. Thus when $a, b \in \mathbb{N}_0$,

$$\gamma_{[y,x]}^{(a,b)} = \left| \left\{ (y_1, \dots, y_a, w_1, \dots, w_b) : 0 \leqslant y_1 \leqslant \dots \leqslant y_a \leqslant y \leqslant w_1 \leqslant \dots \leqslant w_b \leqslant x \right\} \right|.$$

We write $\binom{m}{c}$ for $\binom{m+c-1}{c}$. Note that $\binom{m}{c} = \binom{m+c-1}{c} = \binom{m+c-1}{m-1} = \binom{c+1}{m-1}$, and so, unlike the normal binomial coefficient, $\binom{m}{c}$ is, separately, a polynomial function of both m and c. We use this fact repeatedly below to deduce results on $\gamma^{(a,b)}$ for general $a, b \in \mathbb{R}$ from the special cases when $a, b \in \mathbb{N}_0$.

Lemma 5.1. Let $m, n \in \mathbb{N}$.

- (i) If $c \in \mathbb{N}_0$ then $\sum_{y=0}^{x} {y+1 \choose c} = {x+1 \choose c+1}$. (ii) If $c, d \in \mathbb{N}_0$ then $\sum_{x=0}^{n-1} {n-x \choose c} = {x+1 \choose c+1}$. (iii) If $c, d \in \mathbb{N}_0$ then $\sum_{y=0}^{x} {x-y+1 \choose c} = {x+c+1 \choose c+d+1}$. (iv) If $c \in \mathbb{N}_0$ then $\sum_{y=0}^{x} {x-y+1 \choose c} = {x+c+1 \choose c+d+1}$.

Proof. For (i) we count (c+1)-multisubsets of $\{0,\ldots,x\}$: the (c+1)multisubsets having y as their greatest element are in bijection with the c-multisubsets of $\{0,\ldots,y\}$; these are counted by $\binom{y+1}{c}$. For (ii), use $\binom{n-x}{c} = \binom{c+1}{n-x-1}$ to rewrite the left-hand side as $\sum_{x=0}^{n-1} \binom{c+1}{n-x-1} \binom{d+1}{x}$. This counts the (n-1)-multisubsets of the disjoint union of sets of size c+1 and d+1, and so is $\binom{c+d+2}{n-1} = \binom{n}{c+d+1}$. For (iii), the non-zero summands occur for $y \ge d$. We rewrite $\binom{y}{d}$ as $\binom{y-d+1}{d}$ and apply (ii) to the right-hand side of

$$\sum_{y=d}^{x} \left(\begin{pmatrix} x-y+1 \\ c \end{pmatrix} \right) \left(\begin{pmatrix} y-d+1 \\ d \end{pmatrix} \right) = \sum_{z=0}^{x-d} \left(\begin{pmatrix} x-d+1-z \\ c \end{pmatrix} \right) \left(\begin{pmatrix} z+1 \\ d \end{pmatrix} \right)$$

to get $\binom{x-d+1}{c+d+1}$, which is $\binom{x+c+1}{c+d+1}$, as required. Finally for (iv), we use $\binom{x-y+1}{c} = \binom{x-y+c}{c} = (-1)^{x-y} \binom{-c-1}{x-y}$ to rewrite the right-hand side as $(-1)^x \sum_{y=0}^x \binom{-c-1}{x-y} \binom{m}{y}$; by Vandermonde's convolution this is $(-1)^x \binom{m-c-1}{x}$. as required.

5.2. Stochastic properties of $P(\gamma^{(a,b)})$ and $P(\delta^{(a',b')})$. The function $\gamma^{(a,b)}$ factorizes as the product $\alpha^{(a)}\beta^{(b)}$ where $\alpha^{(a)}$ is defined by $\alpha_y^{(a)} = {y+a \choose y}$ and $\beta^{(b)}$ is defined by $\beta_{[y,x]}^{(b)} = {x-y+b \choose x-y}$. When $a, b \in \mathbb{R}^{>-1}$, $\alpha^{(a)}$ is a strictly positive atomic weight and $\beta^{(b)}$ is a strictly positive \star -symmetric weight. The weight $\delta^{(a',b')} = {a'-1 \choose y}{b'-1 \choose x-y}$ defined for a',b'>1 has a similar factorization. ization. If $b' \notin \mathbb{N}$ then the domain of $\delta^{(a',b')}$ is $\min(\lceil a' \rceil, \lceil b' \rceil)$ and $\delta^{(a',b')}$ is strictly positive. Otherwise $\delta^{(a',b')}$ has the potentially larger domain $\lceil a' \rceil$ and $\delta_{[y,x]}^{(a',b')} > 0$ if and only if y - x < b'. Since $b' \geqslant 2$, the hypothesis of Theorem 1.3 holds. Therefore in all cases the involutive walk is irreducible, recurrent and ergodic with a unique invariant distribution.

To find the invariant distributions we use Theorem 1.4. The following lemma is required.

Lemma 5.2. Let $x, y \in n$ with $y \leq x$. For $a, b \in \mathbb{R}$ we have

(i)
$$\alpha_{[y,x]}^{(a)} = (\binom{y+1}{a})$$
 and $N(\alpha^{(a)})_x = (\binom{x+1}{a+1});$
(ii) $\beta_{[y,x]}^{(b)} = (\binom{x-y+1}{b})$ and $N(\beta^{(b)})_x = (\binom{x+1}{b+1});$
(iii) $\gamma_{[y,x]}^{(a,b)} = (\binom{y+1}{a})(\binom{x-y+1}{b})$ and $N(\gamma^{(a,b)})_x = (\binom{x+1}{a+b+1}).$

Proof. The first parts of (i), (ii) and (iii) are immediate from the definitions. Since $N(\alpha^{(a)})_x$, $N(\beta^{(b)})_x$ and $N(\gamma^{(a,b)})_x$ are polynomials in a and b, it suffices to prove the remaining results when a and b are integral. In this case the normalization factors $N(\alpha^{(a)})_x$ and $N(\beta^{(b)})_x$ follow immediately from Lemma 5.1(i). For $\gamma^{(a,b)}$ we have

$$N(\gamma^{(a,b)})_x = \sum_{y=0}^x \gamma^{(a,b)}_{[y,x]} = \sum_{y=0}^x \left(\! \begin{pmatrix} y+1 \\ a \end{pmatrix} \! \right) \left(\! \begin{pmatrix} x-y+1 \\ b \end{pmatrix} \! \right) = \left(\! \begin{pmatrix} x+1 \\ a+b+1 \end{pmatrix} \! \right)$$

by Lemma 5.1(ii).

Thus $N(\gamma^{(a,b)})_x = {x+a+b+1 \choose x}$ and so $N(\gamma^{(a,b)})_x > 0$ for $x \in n$ if and only if $a+b+2, \ldots, a+b+n \neq 0$, or equivalently, if and only if

$$(5.1) a+b \notin \{-2, \dots, -n\}.$$

We saw after (1.4) in the introduction that $\delta_{[y,x]}^{(a',b')} = (-1)^x \gamma_{[y,x]}^{(-a',-b')}$ for a', $b' \in \mathbb{R}^{>1}$. Hence $N(\delta^{(a',b')})_x = (-1)^x N(\gamma^{(-a,-b)})$. Using that $\binom{m}{c} = (-1)^x N(\gamma^{(-a,-b)})$

 $\binom{c+1}{m-1}$, it follows from Lemma 5.2(iii) that

(5.2)
$$N(\delta^{(a',b')})_x = \binom{(a'-1)+(b'-1)}{x}.$$

Hence as expected from (5.1), $N(\delta^{(a',b')})_x > 0$ whenever $x \in n$ and n is in the domain of $\delta^{(a',b')}$.

By Theorem 1.4, if $a, b \in \mathbb{R}^{>-1}$, the unique invariant distribution for the $\gamma^{(a,b)}$ -weighted involutive walk on n is proportional to $\alpha_{n-1-x}^{(a)}N(\gamma^{(a,b)})_x$. Therefore, by Lemma 5.2, it is π where

(5.3)
$$\pi_x \propto \binom{n-x}{a} \binom{x+1}{a+b+1}.$$

The normalization factor for π is $\binom{n}{2a+b+2}$ by Lemma 5.1(ii). Similarly for $\delta^{(a',b')}$, by (5.2), the unique invariant distribution on n, for n in the domain of $\delta^{(a',b')}$ is ϕ , where

(5.4)
$$\phi_x \propto \binom{a'-1}{x} \binom{(a'-1)+(b'-1)}{x}.$$

Rewriting the first binomial coefficient as $\binom{a'-1}{a'-1-x}$ and applying Vandermonde's convolution shows that when n=a' the normalization factor for ϕ is $\binom{2a'+b'-3}{a'-1}$. In general there is no simple closed form.

5.3. **Spectrum of** $P(\gamma^{(a,b)})$ **and** $P(\delta^{(a',b')})$. To complete the proof of Theorem 1.5, we must show that the eigenvalues of these transition matrices are as claimed. Again we deal with the two families in a uniform way using that

(5.5)
$$H(\gamma^{(-a',-b')}) = H(\delta^{(a',b')}).$$

Whenever (5.1) holds, the eigenvalues of $H(\gamma^{(a,b)})$ are its diagonal entries $\gamma^{(a,b)}_{[x,x}]/N(\gamma^{(a,b)})_x$. By Lemma 5.2(iii), these are $\lambda^{(a,b)}_0,\ldots,\lambda^{(a,b)}_{n-1}$ where

(5.6)
$$\lambda_d^{(a,b)} = \left(\binom{d+1}{a} \right) / \left(\binom{d+1}{a+b+1} \right).$$

The following lemma, analogous to Lemma 4.1, reveals the eigenvectors and, as in §4, verifies the condition in Lemma 2.2(iii) for $H(\gamma^{(a,b)})$ to be a binomial transform. Recall from (2.1) that $v(d) \in \mathbb{R}^n$ is the column vector with entries $v(d)_x = \binom{x}{d}$.

Lemma 5.3. Let $a, b \in \mathbb{R}$ and let $n \in \mathbb{N}$ be such that $a + b \notin \{-2, \dots, -n\}$. Then

$$H(\gamma^{(a,b)})v(d) = \frac{\left(\binom{d+1}{a}\right)}{\left(\binom{d+1}{a+b+1}\right)}v(d).$$

Proof. We have

$$(5.7) \quad \left(H(\gamma^{(a,b)})v(d)\right)_x = \sum_{y=0}^x \left(\!\!\left(\begin{matrix} y+1\\ a\end{matrix}\right)\!\!\right) \left(\!\!\left(\begin{matrix} x-y+1\\ b\end{matrix}\right)\!\!\right) \left(\begin{matrix} y\\ d\end{matrix}\right) / \left(\!\!\left(\begin{matrix} x+1\\ a+b+1\end{matrix}\right)\!\!\right).$$

When $a, b \in \mathbb{N}_0$, by $\binom{y+a}{a}\binom{y}{d} = \binom{a+d}{a}\binom{y+a}{a+d}$ and an instance of Lemma 5.1(iii), the numerator on the right-hand side is

$$\sum_{y=0}^{x} {y+a \choose a} {y \choose d} \left({x-y+1 \choose b} \right) = {a+d \choose a} \sum_{y=0}^{x} {y+a \choose a+d} \left({x-y+1 \choose b} \right)$$
$$= {a+d \choose a} {x+a+b+1 \choose a+b+d+1}.$$

We now use $\binom{x+a+b+1}{a+b+d+1} / \binom{x+1}{a+b+1} = \binom{x+a+b+1}{a+b+d+1} / \binom{x+a+b+1}{a+b+1} = \binom{x}{d} / \binom{a+b+d+1}{a+b+1}$ to find that the right-hand side of (5.7) is

$$\binom{a+d}{a} \binom{x}{d} / \binom{a+b+d+1}{a+b+1} = \binom{x}{d} \binom{d+1}{a} / \binom{d+1}{a+b+1}$$

$$= v(d)_x \binom{d+1}{a} / \binom{d+1}{a+b+1} .$$

This proves the equality claimed in the lemma when $a, b \in \mathbb{N}_0$, and since, by multiplying through by the non-zero quantities $N(\gamma^{(a,b)})_x$ and $\binom{d+1}{a+b+1}$ each side becomes polynomial in a and b, it holds for all claimed a and b. \square

Hence $H(\gamma^{(a,b)})$ is the binomial transform $H^{\lambda^{(a,b)}}$ where $\lambda_d^{(a,b)}$ is as defined in (5.6). By Proposition 2.3(iii), $P(\gamma^{(a,b)})$ and $P(\delta^{(a',b')})$ have the required eigenvalues. This completes the proof of Theorem 1.5 for $\gamma^{(a,b)}$ and $\delta^{(a',b')}$.

Second largest eigenvalue. Since

$$\frac{\lambda_d^{(a,b)}}{\lambda_{d+1}^{(a,b)}} = \frac{\binom{d+a-1}{d-1} / \binom{d+a+b}{d-1}}{\binom{d+a}{d} / \binom{d+a+b+1}{d}} = \frac{d}{d+a} \frac{d+a+b+1}{d} = \frac{a+b+1+d}{a+d}$$

the sequence of eigenvalues of $P(\gamma^{(a,b)})$ is decreasing in absolute value. The same holds for $P(\delta^{(a',b')})$ rewriting the right-hand side as (a'+b'-d-1)/(a'-d). The ratio above also implies that $\lambda_d^{(a,b)} \to 0$ as $d \to \infty$. The rate of convergence of the involutive walk is controlled by the second largest eigenvalue (in absolute value) of $P(\gamma^{(a,b)})$, namely $\lambda_1^{(a,b)} = (a+1)/(a+b+2)$; for $P(\delta^{(a,b)})$, the rewriting (a'-1)/((a'-1)+(b'-1)) is most convenient.

5.4. **Eigenvectors of** $P(\gamma^{(a,b)})$. We now explore the spectral behaviour of $P(\gamma^{(a,b)})$ when $a, b \in \mathbb{R}^{>-1}$ in more detail. By a basic result on reversible transition matrices, the right-eigenvectors of the $n \times n$ matrix $P(\gamma^{(a,b)})$ are orthogonal with respect to the inner product on \mathbb{R}^n defined by

$$\langle v, w \rangle = \sum_{x=0}^{n-1} \pi_x v_x w_x$$

where $\pi_x = \binom{n-x}{a}\binom{x+1}{a+b+1}/\binom{n}{2a+b+2}$ is the invariant distribution in (5.3). For $0 \leq d < n$, let $w(d) \in \mathbb{R}^n$ be the right-eigenvector of $P(\gamma^{(a,b)})$ with eigenvalue $(-1)^d \lambda_d^{(a,b)}$, normalized with respect to $\langle \ , \ \rangle$. By Proposition 2.3(iii), $w(d) \in \langle v(0), \ldots, v(d) \rangle$ for $0 \leq d < n$. Therefore $w(0), \ldots, w(n-1)$ are the

Gram–Schmidt orthonormalizations of $v(0), \ldots, v(n-1)$, with respect to \langle , \rangle . Since the eigenvalues $(-1)^d \lambda_d^{(a,b)}$ of $P(\gamma^{(a,b)})$ are distinct, we also have

$$w(e) \in \left\langle \prod_{d=0}^{e-1} \left(P(\gamma^{(a,b)}) - (-1)^d \lambda_d^{(a,b)} \right) v(e) \right\rangle$$

for $0 \le e < n$. Using either of these observations the right-eigenvectors can be computed for any particular n. In particular, a routine calculation shows that $w(1)_x \propto (a+b+2)(n-1)-(2a+b+3)x$.

The left-eigenvectors can be obtained using the following lemma; we include a proof for completeness and to motivate Definition 8.4 below.

Lemma 5.4. Let P be the transition matrix of a finite reversible Markov chain with strictly positive invariant distribution proportional to π . Let $v \in \mathbb{R}^n$ and let $u \in \mathbb{R}^n$ be defined by $u_x = \pi_x v_x$. Then v is a right-eigenvector for P with eigenvalue λ if and only if u is a left-eigenvector for P with eigenvalue λ .

Proof. We have

$$(uP)_y = \sum_x u_x P_{xy} = \sum_x v_x \pi_x P_{xy} = \sum_x v_x \pi_y P_{yx} = \pi_y \sum_x v_x P_{yx}$$

where the penultimate equality uses the detailed balance equation $\pi_x P_{xy} = \pi_y P_{yx}$. Now v is a right-eigenvector for P with eigenvalue λ if and only if the right-hand side is $\lambda \pi_y v_y$ for each y. Since $u_y = \pi_y v_y$ and $\pi_y > 0$, this holds if and only if u is a left-eigenvector for P with eigenvalue λ .

Remarkably, the left-eigenspace of $P(\gamma^{(a,b)})$ for the final eigenvalue, namely $\lambda_{n-1}^{(a,b)} = (-1)^{n-1} \binom{n}{a} / \binom{n}{a+b+1}$, does not depend on either a or b; up to signs, it is a row of Pascal's Triangle.

Proposition 5.5. Let $u_x = (-1)^x \binom{n-1}{x}$. Then for all $a, b \in \mathbb{R}^{>-1}$ we have $uP(\gamma^{(a,b)}) = (-1)^{n-1} \binom{n}{a} / \binom{n}{a+b+1} u$.

Proof. By Lemma 5.2(iv), $P(\gamma^{(a,b)})_{xy^*} = H(\gamma^{(a,b)})_{xy}$. Hence $\left(uP(\gamma^{(a,b)})\right)_{y^*} = (-1)^{n-1} \binom{n}{a} / \binom{n}{a+b+1} u_{y^*}$ if and only if

$$(5.8) \qquad \sum_{x=y}^{n-1} \frac{\binom{\binom{y+1}{a}}{\binom{\binom{x-y+1}{b}}{\binom{x+1}{a+b+1}}} (-1)^x \binom{n-1}{x} = \frac{\binom{\binom{n}{a}}{\binom{n}{a+b+1}}}{\binom{\binom{n}{a+b+1}}{\binom{n}{a+b+1}}} (-1)^{y^{\star}} \binom{n-1}{y^{\star}}.$$

Observe that when $a, b \in \mathbb{N}$, we have $\binom{n-1}{x}/\binom{x+1}{a+b+1} = \binom{n-1}{x}/\binom{x+a+b+1}{a+b+1} = \binom{n+a+b}{x+a+b+1}/\binom{n+a+b}{a+b+1} = \binom{n+a+b}{n-1-x}/\binom{n}{a+b+1}$. Clearing denominators this becomes a polynomial identity, so it holds for all $a, b \in \mathbb{R}^{>-1}$. Setting w = n - 1 - y, we use this to rewrite the left-hand side as

$$\frac{\binom{\binom{y+1}{a}}{\binom{n}{(a+b+1)}} \sum_{w=0}^{n-1-y} \binom{n-w-y}{b} \binom{n+a+b}{w} (-1)^{n-1-w}}{\binom{n+a+b}{w}} (-1)^{n-1-w}.$$

By an instance of Lemma 5.1(iv), this is $\binom{y+1}{a}(-1)^{n-1-y}\binom{n+a-1}{n-1-y}/\binom{n}{a+b+1}$. Now observe that

$$\binom{y+1}{a}\binom{n+a-1}{n-1-y} = \binom{y+a}{a}\binom{n-1+a}{y+a} = \binom{n}{a}\binom{n-1}{y}$$

which since $\binom{n-1}{y} = \binom{n-1}{y^*}$ gives the relevant factors in the right-hand side of (5.8).

The corollary for the right-eigenvectors of $P(\gamma^{(a,b)})$ is worth noting.

Corollary 5.6. Let $a, b \in \mathbb{R}^{>-1}$. Then

$$w(n-1)_x \propto (-1)^x \binom{n-1}{x} / \binom{n-x}{a} \binom{x+1}{a+b+1}$$
.

Proof. This is immediate from Lemma 5.4 and the invariant distribution in (5.3).

In particular, if a = 0 then from

$$\binom{n+b}{x+b+1} \binom{x+1}{b+1} = \binom{n+b}{x+b+1} \binom{x+b+1}{b+1} = \binom{n+b}{b+1} \binom{n-1}{y}$$

we see that $w(n-1)_x \propto (-1)^x \binom{n+b}{x+b+1}$. A slightly more lengthy argument shows that if a=1 then $w(n-1)_x \propto (-1)^x \binom{n+b+2}{x+b+2}$. However there do not appear to be such 'division-free' formulae for higher a, or general $a \in \mathbb{R}^{>-1}$. We return to these eigenvectors in §9.3.

5.5. The involutive walk on subsets and proof of Corollary 1.10. Let $0 . The <math>\gamma^{(-(1-p),-p)}$ -weighted involutive walk on $\{0,1\}$ has transition matrix

$$Q = \begin{pmatrix} \cdot & 1 \\ p & 1 - p \end{pmatrix}.$$

By special cases of Theorem 1.4 and Theorem 1.5 the invariant distribution is $\left(\frac{p}{1+p}, \frac{1}{1+p}\right)$ and the eigenvalues are 1 and -p. Observe that Q is the transition matrix of the random walk in Corollary 1.10 when m=1; the empty set corresponds to $0 \in \{0,1\}$ and $\{1\}$ to $1 \in \{0,1\}$.

Generally, if P is the matrix of an involutive walk on n and P' is the matrix of an involutive walk on n' then, indexing the entries of $P\otimes P'$ by the ordinal product $n\times n'$, we have $(P\otimes P')_{(x,x'),(z,z')}=P_{(x,z)}P_{(x',z')}$, and so $P\otimes P'$ is the transition matrix of the random walk on $n\times n'$ in which, starting from (x,x'), we choose (y,y') with $y\in [x],y'\in [x']$ with probability $P_{(x,y^*)}P_{(x',y'^*)}$ and then step to (y^*,y'^*) . It is easily seen that if the random walks on n and on n' are irreducible, reversible, recurrent and ergodic then so is the random walk on $n\times n'$. In particular, the involutive walk on subsets of $\{1,\ldots,m\}$ in Corollary 1.10 has all these properties. Its transition matrix is $Q^{\otimes m}$ and so its unique invariant distribution π is $\left(\frac{p}{1+p},\frac{1}{1+p}\right)^{\otimes m}$; by the previous paragraph we have $\pi_X=\left(\frac{p}{1+p}\right)^{m-|X|}\left(\frac{1}{1+p}\right)^{m-|X|}$ for each $X\subseteq\{1,\ldots,m\}$.

Therefore $\pi_X = p^{m-|X|}/(1+p)^m$ as claimed. It is easily seen, either directly or by Proposition 5.5, that the left-eigenvector of Q with eigenvalue -pis (1,-1). Therefore the $(-p)^e$ -eigenspace is spanned by all m-fold tensor products of (p,1) and (1,-1) and has dimension $\binom{m}{e}$. This completes the proof of Corollary 1.10.

6. When is a binomial transform stochastic?

Given a sequence $\lambda_0, \ldots, \lambda_{n-1}$ of real numbers, recall from (1.6) that H^{λ} denotes the binomial transform $B(n)\text{Diag}(\lambda_0,\ldots,\lambda_{n-1})B(n)^{-1}$. By Theorem 2.9, an $n \times n$ matrix has global anti-diagonal eigenbasis action if and only if it is equal to some H^{λ} , and, by Theorem 2.4, this property also characterises generic matrices with the global anti-diagonal eigenvalue property. We have defined $P^{\lambda} = H^{\lambda}J(n)$ to be the anti-triangular matrix corresponding to H^{λ} . In this section we determine when P^{λ} is stochastic, and, supposing that P^{λ} is stochastic, when the involutive walk it defines is ergodic.

The introduction discusses the connection with the results on half-infinite matrices in [6]. Our approach, using the following two lemmas, is more elementary.

Lemma 6.1. Let $\lambda_0, \ldots, \lambda_{n-1} \in \mathbb{R}$ and let $x, y \in n$ with x < n-1. If $x \ge y$

$$H_{x+1,y}^{\lambda} \binom{x+1}{y}^{-1} = H_{xy}^{\lambda} \binom{x}{y}^{-1} - H_{x+1,y+1}^{\lambda} \binom{x+1}{y+1}^{-1}.$$

Proof. Let $k = x - y \in \mathbb{N}_0$. By Lemma 2.2(iv), the coefficient of λ_{y+e} in each of $H_{xy}^{\lambda}\binom{x}{y}^{-1}$, $H_{x+1,y+1}^{\lambda}\binom{x+1}{k}^{-1}$ and $H_{x+1,y}^{\lambda}\binom{x+1}{y}^{-1}$ is $(-1)^e\binom{k}{e}$, $(-1)^{e-1}\binom{k}{e-1}$ and $(-1)^e\binom{k+1}{e}$ respectively, for $e \in \{1,\ldots,n-1-y\}$. For e=0 the coefficient is 1, 0 and 1, respectively. The lemma therefore follows from the basic identity $\binom{k}{e} + \binom{k}{e-1} = \binom{k+1}{e}$ for $e \in \{1, \dots, k\}$.

Lemma 6.2. Let $\lambda_0, \ldots, \lambda_{n-1} \in \mathbb{R}$. Then

- (i) H^{λ} is non-negative if and only if its bottom row is non-negative;
- (ii) if H^{λ} is non-negative and $H^{\lambda}_{xy} = 0$ then $H^{\lambda}_{x'y'} = 0$ for all $x', y' \in n$ with $x' - x \geqslant y' - y$.
 - (iii) $H_{x+1,x}^{\lambda} = {x+1 \choose x} (\lambda_x \lambda_{x+1})$ for all $x \in \{0, 1, \dots, n-2\}$; (iv) if H^{λ} is non-negative and $\lambda_t = \lambda_{t+1}$ then $\lambda_t = \dots = \lambda_{n-1}$.

Proof. As a technical tool, we define F^{λ} to be the lower-triangular $n \times n$ matrix with entries $F_{xy}^{\lambda} = H_{xy}^{\lambda} {x \choose y}^{-1}$ for $x \ge y$ and $F_{xy}^{\lambda} = 0$ for x < y. By Lemma 6.1, the entries of F satisfy the recurrence

$$F_{x+1,y}^{\lambda} = F_{xy}^{\lambda} - F_{x+1,y+1}^{\lambda}.$$

Hence if $F_{xy}^{\lambda} < 0$ then either $F_{x+1,y}^{\lambda} < 0$ or $F_{x+1,y+1}^{\lambda} < 0$. Thus a negative entry in row x of F^{λ} implies a negative entry in row x+1 of F^{λ} . It follows that F^{λ} is non-negative if and only if its bottom row is non-negative, and

the same holds for H^{λ} . This proves (i). Suppose that H^{λ} is non-negative and that $H^{\lambda}_{xy} = 0$. Then $F^{\lambda}_{xy} = 0$ and by the recurrence relation, $F^{\lambda}_{x+1,y} = F^{\lambda}_{x+1,y+1} = 0$. It follows that $H^{\lambda}_{x'y'} = 0$ for all $x' \geqslant x$ and $y' \geqslant y$ with $x' - x \geqslant y' - y$, proving (ii). Part (iii) also follows from the recurrence, using that $F^{\lambda}_{xx} = \lambda_x$. If $\lambda_t = \lambda_{t+1}$ then, by (iii), $F^{\lambda}_{t+1,t} = 0$, and so by (ii), $F^{\lambda}_{x+1,x} = 0$ for all $x \geqslant t$. Hence, by another use of (iii), $\lambda_x = \lambda_{x+1}$ for all $x \geqslant t$, as required.

Corollary 1.6. Let $\lambda_0, \ldots, \lambda_{n-1} \in \mathbb{R}$. The matrix P^{λ} is the transition matrix of an involutive walk if and only if $\lambda_0 = 1$ and

$$\sum_{e=0}^{z} (-1)^e \binom{z}{e} \lambda_{z^*+e} \geqslant 0$$

for all $z \in n$. Moreover if strict inequality holds for all $z \in n$ then $1 > \lambda_1 > \ldots > \lambda_{n-1}$, and $P_{xz}^{\lambda} > 0$ if and only if $x + z \ge n - 1$.

Proof. By Lemma 2.2(iv), $H_{n-1,y}^{\lambda} = \sum_{e=0}^{y} (-1)^e \binom{n-1-y}{e} \lambda_{z^*+e}$. Since this is the entry in column y^* of the bottom row of P^{λ} , it follows from Lemma 6.2(i) that P^{λ} is non-negative if and only if $\sum_{e=0}^{z^*} (-1)^e \binom{z}{e} \lambda_{z^*+e} \geqslant 0$ for all $z \in n$. Moreover, by Lemma 6.2(iv), if strict inequality holds for all $z \in n$ then $\lambda_0 > \lambda_1 > \ldots > \lambda_{n-1}$ and, by Lemma 6.2(ii), P^{λ} has strictly positive entries in all positions below its anti-diagonal. Recall that $v(0) \in \mathbb{R}^n$ is the all-ones vector. By Lemma 2.2(ii), $H^{\lambda}v(0) = \lambda_0 v(0)$. Therefore H^{λ} , and hence P^{λ} , has all row sums equal to 1 if and only if $\lambda_0 = 1$.

Either by taking a linear combination of the inequalities in this corollary, or more directly from Lemma 6.2(iii), it follows that if P^{λ} is a transition matrix then $\lambda_0 \geqslant \lambda_1 \geqslant \ldots \geqslant \lambda_{n-1}$, and so the below-diagonal entries, namely $\binom{x+1}{x}(\lambda_x - \lambda_{x+1})$ for $x \in n$ with $x \leqslant n-2$ are non-negative. Similarly, from the entries below these, we get the convexity property $(\lambda_{x-1} + \lambda_{x+1})/2 \geqslant \lambda_x$ for all $x \in n$ such that $1 \leqslant x < n-1$. As mentioned in the introduction, the full condition in Corollary 1.6 implies that the eigenvalues are completely monotonic. This is of course the case for the eigenvalues $\lambda_d^{(a,b)}$ of the matrices $P(\gamma^{(a,b)})$, $P(\gamma^{(c)})$ and $P(\delta^{(a',b')})$. Since these transition matrices are of irreducible ergodic walks, it is natural to ask, more generally, when is the involutive walk defined by P^{λ} irreducible and ergodic? This has a neat answer, although the proof, using the following proposition, is unavoidably somewhat technical.

Proposition 6.3. Suppose that P^{λ} is the transition matrix of an involutive walk in which 0 is accessible from all states. Let $s \in \mathbb{N}$ be maximal such that $\lambda_{n-s} = \ldots = \lambda_{n-1}$. Then $s \leq n/2$ and $1 > \lambda_1 > \ldots > \lambda_{n-s-1} > \lambda_{n-s} = \ldots = \lambda_{n-1} > 0$ and the walk is irreducible, recurrent and ergodic with a unique strictly positive invariant distribution.

Proof. By Corollary 1.6, $\lambda_0=1$. Since 0 can be reached in one step only from n-1, and $P_{n-1,0}^{\lambda}=H_{n-1,n-1}^{\lambda}=\lambda_{n-1}$, we have $\lambda_{n-1}>0$. Hence, by Lemma 6.2(iii) and (iv), $\lambda_0>\lambda_1>\ldots>\lambda_{n-s-1}>\lambda_{n-s}=\ldots=\lambda_{n-1}>0$. By Lemma 6.2(iii), $H_{n-s+1,n-s}^{\lambda}=\binom{n-s+1}{n-s}(\lambda_{n-s}-\lambda_{n-s+1})=0$. Hence, by Lemma 6.2(ii), $H_{xy}^{\lambda}=0$ for all $x\geqslant n-s+1$, $y\geqslant n-s$ with $x-(n-s+1)\geqslant y-(n-s)$. Thus if y>n-1-s then column y of H^{λ} is zero except for its diagonal entry λ_y . Since $s^{\star}=n-1-s$, it follows that if $y>s^{\star}$ then column y of H^{λ} has a unique non-zero entry of λ_y in row y.

Suppose, for a contradiction, that s > n/2. If n is even, with n = 2m, then $s^* < m - 1$, and the unique non-zero entries of H^{λ} in columns m - 1 and m are on the diagonal. Thus from m-1 the involutive walk must step to $(m-1)^* = m$, and from m it must step to $m^* = m-1$, contradicting that 0 is accessible from any state. Similarly, if n is odd, with n = 2m+1, then $s^* < m$ and from m the involutive walk must step to $m^* = m$. Therefore $s \leq n/2$.

Again by Lemma 6.2(iii), $H_{s^*+1,s^*}^{\lambda} = H_{n-s,n-s-1}^{\lambda} = (n-s)(\lambda_{n-s-1} - \lambda_{n-s}) > 0$. By Lemma 6.1,

$$H_{x,s^\star}^\lambda \binom{x}{s^\star}^{-1} = H_{x-1,s^\star}^\lambda \binom{x-1}{s^\star}^{-1} - H_{x,s^\star+1}^\lambda \binom{x}{s^\star+1}^{-1} = H_{x-1,s^\star}^\lambda \binom{x-1}{s^\star}^{-1}$$

for $x > s^*$. It follows inductively that all the entries on or below the diagonal in column s^* of H^{λ} are non-zero. To summarise so far, we have shown that the down-steps $x \mapsto x$ for $x \in n$, $x \mapsto x - 1$ for $x \in \{1, \dots, s^* + 1\}$ and $n - 1 \mapsto s^*$ have non-zero probability. Moreover, since $s^* + 1 - s = (n - 1 - s) + 1 - s = n - 2s \geqslant 0$, we have $s \leqslant s^* + 1$. It is now routine to check that there is a cycle in the involutive walk

$$n-1 \mapsto s \mapsto s^{\star}+1 \mapsto s-1 \mapsto s^{\star}+2 \mapsto \dots \mapsto n-2 \mapsto 1 \mapsto n-1$$

where bold arrows indicate the steps whose down-step is $x \mapsto x-1$. Observe that t-1 appears two steps before t for each $t \in \{s^*+2, \ldots, n-1\}$. Therefore t-1 is accessible from t for each such t. Since the step $x \mapsto x^*$ has non-zero probability for all x, it follows that $(t-1)^*$ is accessible from t for each $t \in \{s^*+2, \ldots, n-1\}$. Since the step $x \mapsto (x-1)^*$ has positive probability for each $x \in \{1, \ldots, s^*+1\}$ the hypothesis of Lemma 3.1 holds. Therefore the involutive walk is irreducible. Moreover, as in the proof of Theorem 1.3, the step $m \mapsto m$ where $m = \lfloor n/2 \rfloor$ has non-zero probability. Hence the chain is aperiodic and ergodic with a unique strictly positive invariant distribution.

Corollary 6.4. Suppose that P^{λ} is the transition matrix of an involutive walk. Then 0 is accessible from all states if and only if the walk is irreducible, recurrent and ergodic with a unique strictly positive invariant distribution.

Proof. The 'only if' direction is proved by Proposition 6.3. Conversely, if the walk is irreducible then 0 is accessible.

7. BINOMIAL TRANSFORMS AND REVERSIBILITY

Recall that $P^{\lambda} = H^{\lambda}J(n)$ where H^{λ} is the binomial transform defined in (1.6). By Theorem 1.5 and the remark following (1.6), the $\gamma^{(a,b)}$ -, $\delta^{(a',b')}$ - and $\gamma^{(c)}$ -weighted involutive walks on n are reversible for any n in their domain and their transition matrices are of the form P^{λ} . This proves the 'if' direction of Theorem 1.8. The converse is the surprising part and is proved in this section.

Our starting hypothesis is that $n \ge 3$ and P^{λ} is a globally reversible (see Definition 1.7) $n \times n$ transition matrix of an involutive walk in which 0 is accessible from every state. The corresponding down-step matrix H^{λ} is a binomial transform, as in (1.6), with eigenvalues $\lambda_0, \ldots, \lambda_{n-1}$. By Corollary 1.6, $\lambda_0 = 1$. Throughout, we let $\mu = \lambda_1$ and $\nu = \lambda_2$. If $\mu = 1$ then by Lemma 6.2(iii), $H_{10}^{\lambda} = 0$, and so by Lemma 6.2(iii) and (iv) it follows that $P^{\lambda} = J(n)$. This contradicts the assumption that 0 is accessible from every state. Hence

$$(\dagger) 1 > \mu \geqslant \nu \geqslant \lambda_3 \geqslant \dots \geqslant \lambda_{n-1} \geqslant 0.$$

Since the down-step matrices $H(\gamma^{(a,b)})$, $H(\delta^{(a',b')})$ and $H(\gamma^{(c)})$ are binomial transforms, we have the following key principle: if γ is the relevant weight $\gamma^{(a,b)}$, $\gamma^{(c)}$ or $\delta^{(a',b')}$ and $\lambda_d(\gamma) = H(\gamma)_{dd}$ for $d \in \{0, \ldots, n-1\}$ then

(*)
$$P^{\lambda} = P(\gamma) \iff \begin{cases} \text{for all } d \in \{0, \dots, n-1\}, \text{ the eigenvalue } \lambda_d \text{ of } \\ H^{\lambda} \text{ is equal to the eigenvalue } \lambda_d(\gamma) \text{ of } H(\gamma). \end{cases}$$

It is therefore important that the eigenvalues μ and ν determine the parameters a and b. By (1.5), (\star) requires

(7.1)
$$\frac{a+1}{a+b+2} = \mu, \quad \frac{(a+1)(a+2)}{(a+b+2)(a+b+3)} = \nu.$$

Provided $\nu \neq \mu^2$, these equations have the unique solution specified by the functions

(7.2)
$$a(\mu,\nu) = \frac{\mu(\mu-\nu)}{\nu-\mu^2} - 1, \quad b(\mu,\nu) = \frac{(1-\mu)(\mu-\nu)}{\nu-\mu^2} - 1.$$

This gives some motivation for the overall strategy of the proof, which is by induction on n, using the global reversibility hypothesis to show that the three greatest eigenvalues $1, \mu, \nu$ of H^{λ} determine all n eigenvalues $\lambda_0, \ldots, \lambda_{n-1}$, and hence that the right-hand side in (\star) holds.

We structure the argument as follows: §7.1 gives a version of Kolmogorov's Criterion adapted to involutive walks; §7.2 proves Theorem 1.8 when n=3, and motivates §7.3 where we formulate the stronger Proposition 7.3; finally in §7.4 we prove this proposition by induction on n, using Kolmogorov's Criterion on a particular 4-cycle. The case $\nu=2\mu-1$, when this cycle has zero probability, is dealt with separately.

7.1. **Kolmogorov's Criterion.** One form of Kolmogorov's Criterion (see for instance [2, §1.5]) is that a random walk with a strictly positive invariant distribution is reversible if and only if the product of transition probabilities around any cycle does not depend on the direction of travel. The special case for involutive walks is stated in (ii) below.

Lemma 7.1. Let P be the transition matrix of an involutive walk with strictly positive invariant distribution.

- (i) If the walk is reversible then for all $x, z \in n$, we have $P_{xz} > 0$ if and only if $P_{zx} > 0$.
 - (ii) The walk is reversible if and only if

$$P_{x_0x_1}P_{x_1x_2}\dots P_{x_{\ell-1}x_0} = P_{x_0x_{\ell-1}}\dots P_{x_2x_1}P_{x_1x_0}$$

for all distinct $x_0, x_1, \ldots, x_{\ell-1} \in n$ with $\ell \geq 3$, such that $x_i + x_{i+1} \geq n-1$ for all $i \in \{0, 1, \ldots, \ell-2\}$, and $x_{\ell-1} + x_0 \geq n-1$.

Proof. From the detailed balance equations $\pi_x P_{xz} = \pi_z P_{zx}$ we get $P_{xz} > 0$ if and only if $P_{zx} > 0$, proving (i). Since $P_{xz} = 0$ unless $x + z \ge n - 1$, it suffices to verify Kolmogorov's Criterion for cycles of the form $x_0 \mapsto x_1 \mapsto x_2 \mapsto \ldots \mapsto x_{\ell-1} \mapsto x_0$ where $x_i + x_{i+1} \ge n - 1$ for all $i \in \{0, 1, \ldots, \ell - 2\}$ and $x_{\ell-1} + x_0 \ge n - 1$. This is trivial for 2-cycles, so we may suppose that $\ell \ge 3$, giving (ii).

7.2. **Base case:** n=3. If $\nu>\mu^2$ then defining a and b by (7.2) we have $a, b \in \mathbb{R}^{>-1}$ and the eigenvalues of H^{λ} and $H(\gamma^{(a,b)})$ agree, so by (\star) , $H^{\lambda}=H(\gamma^{(a,b)})$. If $\nu=\mu^2$ then $\mu<1$ by (\dagger) , and so by (\star) , $H^{\lambda}=H(\gamma^{(c)})$ where $1/(c+1)=\mu$, or equivalently, $c=\frac{1-\mu}{\mu}$. In the remaining case $\nu<\mu^2$ and again defining a and b by (7.2) we have a, b<-1. By either Lemma 2.2(iv) or Lemma 6.1,

$$H^{\lambda} = \begin{pmatrix} 1 & \cdot & \cdot \\ 1 - \mu & \mu & \cdot \\ 1 - 2\mu + \nu & 2(\mu - \nu) & \nu \end{pmatrix}$$

and so $\nu \ge 2\mu - 1$. Moreover, since 0 is accessible from all states, $\nu > 0$. Setting $\nu = 2\mu - 1 + \tau(1 - \mu)$ we have $\mu - \nu = (1 - \mu)(1 - \tau)$ and $\mu - \nu^2 = (1 - \mu)(1 - \mu - \tau)$, hence $(1 - \mu)(\mu - \nu)/(\nu - \mu^2) = -(1 - \mu)(1 - \tau)/(1 - \mu - \tau)$ where $1 - \mu - \tau > 0$ and

$$-b(\mu, 2\mu - 1 + \tau(1-\mu)) = 2 + \frac{\mu\tau}{1-\mu-\tau},$$

for $0 \le \tau < 1 - \mu$. Similar calculations show that

$$-a(\mu, 2\mu - 1 + \tau(1-\mu)) = \frac{1 - \tau - \mu\tau}{1 - \tau - \mu}.$$

Comparing the numerator and denominator immediately above we see that $-a(\mu, 2\mu - 1 + \tau(1-\mu))$ is an increasing function of τ with minimum $\frac{1}{1-\mu}$

when $\tau = 0$. Since $\nu > 0$ we have $\mu > \frac{1}{2}$ and so a' > 2. Moreover, when $\tau = 0$, we have b' = 2 as in the final case of Theorem 1.8. Otherwise $\tau > 0$ and b' > 2, again as required. This completes the proof of the 'only if' direction when n = 3.

7.3. **Inductive claim.** To enable an inductive proof, we must prove a stronger version of Theorem 1.8 that takes into account that the exceptional case $\nu = 2\mu - 1$ and $a' = \frac{1}{1-\mu}$ is the first in an infinite family.

To motivate the statement of Proposition 7.3, observe that if $m \in \mathbb{N}$ then $\delta_{[y,x]}^{(a',m)} = {a'-1 \choose y} {m-1 \choose x-y}$ vanishes if and only if $x-y \geqslant m$. Thus provided $n \geqslant m$, the transition matrix $P(\delta^{(a',m)})$ has exactly m non-zero anti-diagonal bands. This was seen, when m=2, in one of the examples before Theorem 1.5, and also in the base case just proved, for which $-b(\mu,\nu) = b(\mu,2\mu-1) = 2$ and $\mu > \frac{1}{2}$. Therefore we may assume $\mu > 1/2$ when finding these exceptional cases. Let $b'_{\mu}(\nu) = -b(\mu,\nu)$ and note that

$$\frac{db'_{\mu}(\nu)}{d\nu} = \frac{\mu(1-\mu)^2}{(\mu^2 - \nu)^2}.$$

Thus b'_{μ} is increasing on the interval $[0, \mu^2)$, with $b'_{\mu}(0) = \frac{1}{\mu}$ and $b'_{\mu}(\nu) \to \infty$ as $\nu \to \mu^2$. Since $\mu > \frac{1}{2}$, for each $m \in \mathbb{N}$ with $m \geqslant 2$ there exists a unique ν such that $b'_{\mu}(\nu) = m$. We denote this value by $\nu_m(\mu)$. Let $a'_m(\nu) = -a(\mu, \nu_m(\mu))$. By inverting b'_{μ} and then substituting one gets the explicit formulae

(7.3)
$$\nu_m(\mu) = \frac{\mu(m\mu - 1)}{m - 2 + \mu}, \quad a'_m(\mu) = \frac{(m - 2)\mu + 1}{1 - \mu}.$$

In particular, $\nu_2(\mu) = 2\mu - 1$ and $a_2'(\mu) = \frac{1}{1-\mu}$, as seen in the case n = 3. Since b_{μ}' is an increasing function, so is the sequence $\nu_m(\mu)$. Moreover, since $b_{\mu}'(\nu) \to \infty$ as $\nu \to \mu^2$, the sequence $\nu_m(\mu)$ converges to μ^2 as $m \to \infty$. This can be seen in the graph in Figure 1 showing $a(\frac{2}{3}, \nu)$ and $b(\frac{2}{3}, \nu)$ and the exceptional values $\nu_m(\frac{2}{3})$ for $m \in \{2, 3, 4, 5, 6\}$.

Instead fixing m, the relevant values of n are given the lemma below, which shows that the chosen domains of $\delta^{(a',b')}$ for $b' \notin \mathbb{N}$ and $\delta^{(a',b')}$ for $b' \in \mathbb{N}$ specified after (1.4) are as large as possible.

Lemma 7.2. Let $a', b' \in \mathbb{R}^{>1}$.

- (i) If $b' \notin \mathbb{N}$ then the function $\delta^{(a',m)}$ on intervals of n is a weight such that 0 is accessible from every state in the $\delta^{(a',b')}$ -weighted involutive walk if and only if $n \leq \min(\lceil a' \rceil, \lceil b' \rceil)$ and no larger domain is possible.
- (ii) Suppose that $b' = m \in \mathbb{N}$. The function $\delta^{(a',m)}$ on intervals of n is a weight such that 0 is accessible from every state in the $\delta^{(a',m)}$ -weighted involutive walk if and only if $n \leq \lceil a' \rceil$. Moreover, $m \geq \lfloor \frac{1-\mu}{\mu}(n-2) \rfloor + 2$, where μ is the usual eigenvalue.

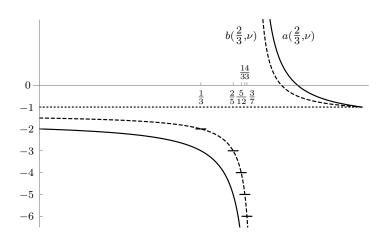


FIGURE 1. The functions $a(\frac{2}{3},\nu)$ and $b(\frac{2}{3},\nu)$ for $0 \le \nu \le 2/3$ and the exceptional values $\nu_m(\frac{2}{3})$ for $m \in \{2,3,4,5,6\}$, namely $\frac{1}{3}, \frac{2}{5}, \frac{4}{12}, \frac{14}{33}, \frac{3}{7}$. Each function has a vertical asymptote at $(\frac{2}{3})^2$.

Proof. By (1.4), $\delta_{[y,x]}^{(a',m)} = {a'-1 \choose y} {b'-1 \choose x-y}$. Thus if $n > \lceil a' \rceil$ then $\delta_{[n-1,n-1]}^{(a',b')}$ is either negative, contrary to the definition of weight, or zero, implying that $(n-1)^\star = 0$ is inaccessible from n-1, and so inaccessible from any state. Similarly, if $b' \not\in \mathbb{N}$ then taking $n > \lceil b' \rceil$ we find that $\delta_{[0,n-1]}^{(a',b')}$ is negative. Thus the largest possible domain is $\min(\lceil a' \rceil, \lceil b' \rceil)$, as claimed. When $b' \in \mathbb{N}$, as in (ii), there is no constraint from ${b'-1 \choose x-y}$, and so the largest domain is instead $\lceil a' \rceil$. To complete the proof, we use the equation for $a'_m(\mu)$ in (7.3) to get $n \leqslant \lceil \frac{(m-2)\mu+1}{1-\mu} \rceil$. Hence $n < \frac{(m-2)\mu+1}{1-\mu} + 1$, and so $(1-\mu)n < (m-2)\mu+2-\mu$ which rearranges to $m > \frac{1-\mu}{\mu}n-\frac{2}{\mu}+3=\frac{1-\mu}{\mu}(n-2)+1$, or equivalently, $m \geqslant \lfloor \frac{1-\mu}{\mu}(n-2)\rfloor +2$.

We can now state our precise result, knowing that all the transition matrices specified are well-defined and give involutive walks in which 0 is accessible from every state; the condition $\mu > \frac{1}{2}$ in (iv) was seen when n=3 in the proof of the base case.

Proposition 7.3. Let P^{λ} be the transition matrix of a reversible Markov chain on n where $n \geq 3$ in which 0 is accessible from every state. Then $\lambda_0 = 1$. Let $\mu = \lambda_1$, $\nu = \lambda_2$. One of the following cases applies.

- (i) If $\nu > \mu^2$ then $\mu > \nu$ and the unique $a, b \in \mathbb{R}^{>-1}$ such that $P^{\lambda} = P(\gamma^{(a,b)})$ are $a(\mu,\nu)$ and $b(\mu,\nu)$.
 - (ii) If $\nu = \mu^2$ then $\mu < 1$ and $P^{\lambda} = P(\gamma^{(c)})$ where $c = (1 \mu)/\mu$.
- (iii) If $\nu_{n-1}(\mu) < \nu < \mu^2$ then the unique a', $b' \in \mathbb{R}^{>1}$ such that $P^{\lambda} = P(\delta^{(a',b')})$ are $-a(\mu,\nu)$ and $-b(\mu,\nu)$ and $n \leq \min(\lceil a' \rceil, \lceil b' \rceil)$.
- (iv) Otherwise, $\mu > \frac{1}{2}$, $\nu = \nu_m(\mu)$ for a unique $m \in n$ such that $m \geqslant \lfloor \frac{1-\mu}{\mu}(n-2) \rfloor + 2$ and the unique $a' \in \mathbb{R}^{>1}$ such that $P^{\lambda} = P(\delta^{(a',m)})$ is

 $a'_m(\nu)$. Moreover $n \leqslant \lceil a'_m(\nu) \rceil$ and $P(\delta^{(a',m)})$ has exactly m anti-diagonal bands where m < n.

The example below shows the interesting behaviour in (iii) and (iv).

Example 7.4. Take n=10. As in Figure 1, we take $\mu=\frac{2}{3}$. A special feature of this case is that $a'_m(\mu)=2m-1$ is integral. By (iii) and (iv) in Proposition 7.3, either $\nu>\nu_9(\frac{2}{3})=\frac{10}{23}$, or ν , a', b' are in the table below.

By (iv), or equivalently Lemma 7.2, we require $m \geqslant \lfloor \frac{1-\mu}{\mu}(n-2) \rfloor + 2 = \frac{n}{2} + 1 = 6$, and so the table stops with m=6. The next entry, were it permitted, would be $\nu = \frac{14}{45}$, a'=9, b'=5, but $9 \not > 10-1$, and correspondingly 10 is not in the domain of the weight $\delta^{(9,5)}$. Thus when n=10, the reversible random walks with second largest eigenvalue $\frac{2}{3}$ classified in Proposition 7.3 split into a continuously varying family when $\frac{4}{9} < \nu < \frac{2}{3}$, a one-off case when $\nu = \frac{4}{9}$, a continuously varying family when $\frac{10}{23} < \nu < \frac{4}{9}$, and final special cases for $\nu = \frac{10}{23}, \frac{13}{30}, \frac{22}{51}, \frac{3}{7}$, in which the transition matrix has respectively 9, 8, 7 and 6 non-zero anti-diagonal bands. To illustrative the inductive step in the proof of Proposition 7.3 suppose we increase n to 11. Then we require either $\nu > \nu_{10}(\frac{2}{3}) = \frac{17}{39}$ or $\nu = \frac{17}{39}$, a'=19, m=10, losing the cases where $\frac{10}{23} < \nu < \frac{17}{39}$ for n=10.

7.4. Proof of Proposition 7.3.

Case $\nu = 2\mu - 1$. In this case P^{λ} has just two non-zero anti-diagonals. We exploit this in the proof.

Lemma 7.5. Let $n \geqslant 3$. If $\nu = 2\mu - 1$ then $H^{\lambda} = H(\delta^{(\frac{1}{1-\mu},2)})$ and (iv) in Proposition 7.3 holds with m = 2.

Proof. It follows either directly from (1.4) or using Lemma 6.1, that the matrix $H(\delta^{(\frac{1}{1-\mu},2)})$ has entries

(7.4)
$$H(\delta^{(\frac{1}{1-\mu},2)})_{xy} = \begin{cases} x\mu - (x-1) & \text{if } x = y\\ x(1-\mu) & \text{if } x = y+1\\ 0 & \text{otherwise.} \end{cases}$$

When n=3, this has the same eigenvalues, namely 1, μ , $2\mu-1$ as H^{λ} , and so the matrices are equal by (\star) , and by the base case, $\mu>\frac{1}{2}$. Suppose inductively that (\star) holds for n-1. Thus the $n\times n$ matrix P^{λ} has topright $(n-1)\times(n-1)$ submatrix $P(\delta^{(\frac{1}{1-\mu},2)})$. In particular, $\lambda_x=x\mu-(x-1)$ for $x\in n-1$ and, by (7.4), we have $H^{\lambda}_{n-3,0}=H(\delta^{(\frac{1}{1-\mu},2)})_{n-3,0}=0$.

Hence $P_{n-3,n-1}^{\lambda}=0$ and using the reversibility hypothesis, it follows from Lemma 7.1(i) that $P_{n-1,n-3}^{\lambda}=0$. But by Lemma 2.2(iv),

$$P_{n-1,n-3}^{\lambda} = \binom{n-1}{n-3} (\lambda_{n-3} - 2\lambda_{n-2} + \lambda_{n-1}).$$

Therefore

$$((n-3)\mu - (n-4)) - 2((n-2)\mu - (n-3)) + \lambda_{n-1} = 0$$

which simplifies to $-(n-1)\mu + (n-2) + \lambda_{n-1} = 0$. Hence $\lambda_{n-1} = (n-1)\mu - (n-2)$ which is $\lambda_{n-1}^{(\frac{1}{1-\delta},2)}$ by (7.4). Therefore by (\star) , $P^{\lambda} = P(\delta^{(\frac{1}{1-\mu},2)})$. Thus $a' = \frac{1}{1-\mu} = a'_2(\nu)$. Moreover, by Lemma 7.2, $2 > \lfloor \frac{1-\mu}{\mu}(n-2) \rfloor + 2$. Hence $\frac{1-\mu}{\mu}(n-2) < 1$ which implies $n-1 < \frac{1}{1-\mu}$, hence $n \leq \lceil a'_2(\nu) \rceil$, giving the final requirement for (iv).

Inductive step. Finally we are ready to prove Proposition 7.3. Suppose that $n \ge 4$. By induction and the global reversibility hypothesis, there is a weight γ of the required binomial type such that H^{λ} agrees with $H(\gamma)$ except perhaps in its bottom row. Moreover, the parameters for γ (namely a and b when $\gamma = \gamma^{(a,b)}$, c when $\gamma = \gamma^{(c)}$ and a', b' when $\gamma = \delta^{(a',b')}$) are unique. Restricting further to the top-right 3×3 submatrix, the base case implies that $\mu > \nu \ge 2\mu - 1$. If $\nu = 2\mu - 1$ then we are done by Lemma 7.5. Therefore we may assume that $1 > \mu > \nu > 2\mu - 1$.

Consider the 4-cycle

$$n-1 \mapsto 1 \mapsto n-2 \mapsto 2 \mapsto n-1.$$

(If n=4 then instead take the 3-cycle $3\mapsto 1\mapsto 2\mapsto 3$ obtained by deleting the repeated vertex 2.) By the version of Kolmogorov's Criterion in Lemma 7.1(ii), and the hypothesis that P^{λ} is reversible, we have

$$(7.5) P_{n-1,1}^{\lambda} P_{1,n-2}^{\lambda} P_{n-2,2}^{\lambda} P_{2,n-1}^{\lambda} = P_{n-1,2}^{\lambda} P_{2,n-2}^{\lambda} P_{n-2,1}^{\lambda} P_{1,n-1}^{\lambda}.$$

(This holds when n=4 since we have simply introduced an extra term of $P_{2,2}^{\lambda}$.) By the inductive assumption $P_{xz}^{\lambda} = P(\gamma)_{xz}$ for all $x \in n-1$. Moreover by Theorem 1.5 each γ -weighted involutive walk is reversible, and so (7.5) holds when every P_{xz}^{λ} is replaced with $P(\gamma)_{xz}$. That is,

(7.6)
$$P(\gamma)_{n-1,1}P(\gamma)_{1,n-2}P(\gamma)_{n-2,2}P(\gamma)_{2,n-1} = P(\gamma)_{n-1,2}P(\gamma)_{2,n-2}P(\gamma)_{n-2,1}P(\gamma)_{1,n-1}.$$

Let q be the common value in (7.6). Note that $q \neq 0$, since $\mu > 2\nu - 1$ and so γ is not the exceptional weight $\delta^{(\frac{1}{1-\mu},2)}$, the matrix $P(\gamma)$ has at least 3 non-zero anti-diagonals and all terms in the product are non-zero. Thus

$$\begin{split} P_{1,n-2}^{\lambda}P_{n-2,2}^{\lambda}P_{2,n-1}^{\lambda} &= P(\gamma)_{1,n-2}P(\gamma)_{n-2,2}P(\gamma)_{2,n-1} = q/P(\gamma)_{n-1,1} \\ P_{2,n-2}^{\lambda}P_{n-2,1}^{\lambda}P_{1,n-1}^{\lambda} &= P(\gamma)_{2,n-2}P(\gamma)_{n-2,1}P(\gamma)_{1,n-1} = q/P(\gamma)_{n-1,2}. \end{split}$$

By Lemma 2.2(iv), we have $P_{n-1,1}^{\lambda} = \binom{n-1}{1}(\lambda_{n-2} - \lambda_{n-1})$ and $P_{n-1,2}^{\lambda} = \binom{n-1}{2}(\lambda_{n-3} - 2\lambda_{n-2} + \lambda_{n-1})$. Again the analogous equation holds for $P(\gamma)$. Moreover, by the induction hypothesis and (\star) , $\lambda_x = H_{xx}^{\lambda} = H(\gamma)_{xx} = \lambda_x(\gamma)$ for $x \in n-1$. Therefore we obtain another pair of equations

$$P_{n-1,1}^{\lambda} = P(\gamma)_{n-1,1} - \binom{n-1}{1} (\lambda_{n-1} - \lambda_{n-1}(\gamma))$$

$$P_{n-1,2}^{\lambda} = P(\gamma)_{n-1,2} + \binom{n-1}{2} (\lambda_{n-1} - \lambda_{n-1}(\gamma)).$$

Writing Δ for $\lambda_{n-1} - \lambda_{n-1}(\gamma)$ and making the substitutions in (7.5) indicated by the four displayed equations above we obtain

$$(P(\gamma)_{n-1,1} - (n-1)\Delta)\frac{q}{P(\gamma)_{n-1,1}} = (P(\gamma)_{n-1,2} + \binom{n-1}{2}\Delta)\frac{q}{P(\gamma)_{n-1,2}}.$$

Using that $P(\gamma)_{n-1,z} = \gamma_{[z^*,n-1]}/N(\gamma)_{n-1}$ and cancelling all common terms, including $N(\gamma)_{n-1}$, we obtain

$$-(n-1)\frac{\Delta}{\gamma_{[n-2,n-1]}} = \binom{n-1}{2} \frac{\Delta}{\gamma_{[n-3,n-1]}}$$

or equivalently

(7.7)
$$((n-1)\gamma_{[n-3,n-1]} + \binom{n-1}{2}\gamma_{[n-2,n-1]})\Delta = 0.$$

If γ is the weight $\gamma^{(c)}$ then $\gamma^{(c)}_{[y,n-1]} = \binom{n-1}{y}c^{n-1-y}$ and so (7.7) becomes $(n-1)\binom{n-1}{2}c(c-1)\Delta=0$. Since c>1, we have $\Delta=0$. Hence $\lambda_{n-1}=\lambda(\gamma^{(c)})_{n-1}$ and by (\star) we get (ii) in Proposition 7.3. If γ is the weight $\gamma^{(a,b)}$ then $\gamma^{(a,b)}_{[y,n-1]} = \binom{a+y}{y}\binom{b+n-1-y}{n-1-y}$ and so (7.7) becomes

$$\left((n-1) \binom{a+n-3}{n-3} \binom{b+2}{2} + \binom{n-1}{2} \binom{a+n-2}{n-2} \binom{b+1}{1} \right) \Delta = 0.$$

Using $\binom{m+1}{c+1} = \binom{m}{c} \frac{m+1}{c+1}$ this simplifies to

$$\frac{1}{2} \left((n-1) \binom{a+n-3}{n-3} \binom{b+1}{1} \left(\frac{b+2}{2} + \frac{n-2}{2} \frac{a+n-2}{n-2} \right) \Delta = 0,$$

and hence to $(a+b+n)\Delta = 0$. Since $a, b \in \mathbb{R}^{>-1}$ and $n \ge 4$, again we get $\Delta = 0$, and by (\star) we get (i). Finally suppose that γ is the weight $\delta^{(a',b')}$ and so $\gamma_{[y,n-1]} = {a'-1 \choose y} {b'-1 \choose n-1-y}$. Substituting in (7.7) we get

$$\left((n-1) \binom{a'-1}{n-3} \binom{b'-1}{2} + \binom{n-1}{2} \binom{a'-1}{n-2} \binom{b'-1}{1} \right) \Delta = 0$$

which simplifies by $\binom{m}{c+1} = \binom{m}{c} \frac{m-c}{c+1}$ to

$$(n-1)\binom{a'-1}{n-3}\binom{b'-1}{1}\left(\frac{b'-2}{2} + \frac{n-2}{2}\frac{a'-n+2}{n-2}\right)\Delta = 0$$

and hence to $(a'+b'-n)\Delta=0$. In either inductive case we have $a'\geqslant n-1$ and $b'\geqslant 2$, hence a'+b'-n>0 and so $\Delta=0$. By (\star) we get $P^\lambda=P(\delta^{(a',b')})$.

Since 0 is accessible from every state, Lemma 7.2 implies that $n\leqslant \lceil a'\rceil$. If the inductive case is (iv) then (iv) still holds: $\mu>\frac{1}{2},\ \nu=\nu_m(\mu)$ and $a'=a'_m(\mu)$ are inductive assumptions, and $m\geqslant \lfloor\frac{1-\mu}{\mu}(n-2)\rfloor+2$ and $n\leqslant \lceil a'_m(\nu)\rceil$ follow from Lemma 7.2. Finally suppose the inductive case is (iii), so $\nu_{n-2}(\mu)<\nu<\mu^2$. If we have the stronger inequality $\nu_{n-1}(\mu)<\nu$ then (iii) still holds. In the remaining case (iii) holds with $\nu_{n-2}(\mu)<\nu\leqslant\nu_{n-1}(\mu)$. Since b'_μ is an increasing function of ν , we have $n-2< b'\leqslant n-1$. If n-2< b'< n-1 then $\delta^{(a',b')}_{[0,n-1]}=\binom{b'-1}{n-1}<0$ contradicts that $P(\delta^{(a',b')})$ is equal to the stochastic matrix P^λ . Therefore $b'=n-1,\ \nu=\nu_{n-1}(\mu)$ and $a'=a'_{n-1}(\mu)$. (This is the case where $P(\delta^{(a'_{n-1}(\mu),n-1)})$ has n-1 non-zero anti-diagonal bands and a zero in its bottom right corner.) By Lemma 7.2, $n-1>\frac{1-\mu}{\mu}n+2-\frac{1}{\mu}$. If $\mu\leqslant\frac{1}{2}$ then the right-hand side is at least n, a contradiction. Therefore $\mu>\frac{1}{2}$ and we have shown that (iv) holds. This completes the proof.

8. The involutive walk on [0,1]

We now turn to involutive walks defined on the real interval $[0,1] = \{x \in \mathbb{R} : 0 \leq x \leq 1\}$ with its usual order and anti-involution $x^* = 1 - x$. From now on [y,x] denotes a subinterval of [0,1].

Definition 8.1. A real weight is a function γ on the set of non-empty intervals of [0,1] taking values in $\mathbb{R}^{\geqslant 0}$ such that, for each $x \in [0,1]$, the function $y \mapsto \gamma_{[y,x]}$ is continuous and non-zero almost everywhere in [0,x], and $\int_0^x \gamma_{[y,x]} dy < \infty$.

Thus a real weight defines a probability distribution on each interval [0, x] having probability density function $\gamma_{[y,x]}/N(\gamma)_x$ where $N(\gamma)_x = \int_0^x \gamma_{[y,x]} \, \mathrm{d}y$. The positivity assumption in the definition makes these weights the analogue of the strictly positive discrete weights in Definition 1.1. We define atomic and \star -symmetric for real weights as in this definition. The γ -weighted involutive walk on [0,1] is defined as in Definition 1.2, replacing $\{0,1,\ldots,n-1\}$ with [0,1].

8.1. Invariant distributions.

Proposition 8.2. Let α be an atomic real weight and let β be a \star -symmetric real weight. Suppose that $\int_0^1 \alpha_{1-x} N(\alpha\beta)_x dx < \infty$. Then the $\alpha\beta$ -weighted involutive walk on n is reversible with respect to an invariant distribution proportional to $\alpha_{1-x} N(\alpha\beta)_x$.

Proof. It is routine to solve the detailed balance equation using essentially the same arguments used to prove Lemma 3.2 and Lemma 3.3.

The integrability assumption in the lemma is essential.

Example 8.3. Let $a \in \mathbb{R}$ with a > 0. Let $\alpha_{[y,x]} = 1/y^a(1-y)^{a+1}$ and let $\beta_{[y,x]} = y^a(1-x)^a$ for $0 < y \le x$. Observe that α and β are real weights and that α is atomic and β is \star -symmetric. Let $\gamma = \alpha\beta$, so $\gamma_{[y,x]} = (1-x)^a/(1-y)^{a+1}$. We have

$$N(\gamma)_x = \int_0^x \frac{(1-x)^a}{(1-y)^{a+1}} \, \mathrm{d}y = \frac{1}{a} (1 - (1-x)^a)$$

hence γ is a real weight. However

$$\alpha_{1-x}N(\gamma)_x = \frac{1}{ax^{a+1}} \left(\frac{1}{(1-x)^a} - 1\right)$$

is not integrable on any interval containing 1, and not integrable on any interval containing 0 whenever $a \ge 1$.

8.2. **Hilbert spaces.** In this section we describe a setting in which the analogues of left- and right-multiplication by the transition matrix for a discrete involutive walk are compact linear operators on Hilbert space. Our §5.4 and [1] suggest the correct inner-product spaces in which to work. Let α be an atomic real weight and let β be a *-symmetric real weight. Define $K:[0,1]^2 \to \mathbb{R}$ by

(8.1)
$$K(x,z) = [x+z \ge 1] \frac{\sqrt{\alpha_{1-z}}\sqrt{\alpha_{1-z}}\beta_{[1-z,x]}}{\sqrt{N(\alpha\beta)_z}\sqrt{N(\alpha\beta)_x}}$$

where $[x+z \ge 1]$ is an Iverson bracket. When $K(x,z) \in L^2([0,1]^2)$, the integral operator $M: L^2([0,1]) \to L^2([0,1])$ defined by

$$(Mf)(x) = \int_0^1 K(x, z) f(z) \, \mathrm{d}z$$

is compact. (See for instance [7, Theorem 8.8].) Since K(x, z) = K(z, x) for all $x, z \in [0, 1]$, M is also self-adjoint.

Let $\vartheta:[0,1]\to\mathbb{R}^{\geqslant 0}$ be the function proportional to the invariant distribution given by Proposition 8.2, as defined by $\vartheta_x=\alpha_{1-x}N(\alpha\beta)_x$. Note that, by Definition 8.1, ϑ_x is non-zero almost everywhere. Given Example 8.3, we also assume that $\sqrt{\vartheta_x}$ is in the Hilbert space $L^2[0,1]$ of square-integrable functions on [0,1].

Definition 8.4. Let $\mathcal{H}_L = \{f(x)/\sqrt{\vartheta_x} : f \in L^2([0,1])\}$ with inner product $\langle , \rangle_{\mathcal{H}_L}$ defined by $\langle f, g \rangle = \int_0^1 \vartheta_x f(x) g(x) \, \mathrm{d}x$. Let $\mathcal{H}_R = \{f(x)\sqrt{\vartheta_x} : f \in L^2([0,1])\}$ with inner product $\langle , \rangle_{\mathcal{H}_R}$ defined by $\langle f, g \rangle = \int_0^1 f(x) g(x)/\vartheta_x \, \mathrm{d}x$.

The maps sending $f \in L^2([0,1])$ to $f/\sqrt{\vartheta_x} \in \mathcal{H}_L$ and $f\sqrt{\vartheta_x} \in \mathcal{H}_R$ are isometric isomorphisms from $L^2([0,1])$ to \mathcal{H}_L and \mathcal{H}_R , respectively. (The first map is well-defined because the zeros of ϑ_x form a null-set.) Therefore \mathcal{H}_L and \mathcal{H}_R are Hilbert spaces. Moreover the constant function is in \mathcal{H}_L and, by our integrability assumption, ϑ_x is in \mathcal{H}_R . The analogues of left-

and right-multiplication by the transition matrix are the integral operators $L_P: \mathcal{H}_L \to \mathcal{H}_L$ and $R_P: \mathcal{H}_R \to \mathcal{H}_R$ defined by

$$(L_P g)(x) = \frac{M(g(z)\sqrt{\vartheta_z})}{\sqrt{\vartheta_x}} \quad \text{for } g \in \mathcal{H}_L$$
$$(R_P g)(z) = M(\frac{g(x)}{\sqrt{\vartheta_x}})\sqrt{\vartheta_z} \quad \text{for } g \in \mathcal{H}_R.$$

(Note we deliberately use different variables: for left-multiplication we integrate over the final state z; for right-multiplication we integrate over the initial state x.) Equivalently,

(8.2)
$$(L_P g)(x) = \int_{1-x}^{1} \frac{\alpha_{1-z} \beta_{[1-z,x]}}{N(\alpha \beta)_x} g(z) dz$$

(8.3)
$$(R_P g)(z) = \int_{1-z}^1 \frac{\alpha_{1-z} \beta_{[1-z,x]}}{N(\alpha \beta)_x} g(x) dx.$$

Since L_P and R_P are conjugate by isometric isomorphisms to the self-adjoint compact operator M, they are also self-adjoint and compact, and have the same eigenvalues.

We can now give a convenient sufficient condition for the real involutive walk to converge. The proof uses the generalization of the Perron–Frobenius Theorem to operators on Banach spaces. Let $\pi_x = A\vartheta_x$ be the invariant probability distribution obtained by normalizing ϑ .

Proposition 8.5. The invariant distribution π_x is unique and the $\alpha\beta$ -weighted involutive walk started at any probability distribution $\vartheta \in \mathcal{H}_R$ converges to $\pi_x \in \mathcal{H}_R$.

Proof. The condition in Definition 8.1 that $\gamma_{[y,x]} > 0$ for almost all $y \in [0,x]$ implies that if \mathcal{C} is the cone in \mathcal{H}_R of non-negative functions then $R_P : \mathcal{C} \setminus \{0\} \to \mathcal{C}^\circ$, where \mathcal{C}° is the interior of C. Hence, by the Krein–Rutman Theorem (see [3, Theorem 6.3]), 1 is the unique largest eigenvalue (in modulus) of R_P , and π_x is the unique 1-eigenfunction. Since R_P is compact and self-adjoint, there is a complete orthonormal basis of \mathcal{H}_R of eigenfunctions; expanding ϑ in this basis and then applying R_P^t shows that $R_P^t \vartheta$ converges, in \mathcal{H}_R , to the invariant distribution π , as $t \to \infty$.

8.3. **Trigonometric example.** As an illustrative example, to which we return in the final subsection, we take $\alpha_x = \sin \pi x$ for $x \in [0,1]$ and $\beta_{[y,x]} = 1$ for each non-empty interval [y,x]. Note that α and β are real weights, α is atomic and β is symmetric. We have $N(\alpha)_x = \int_0^x \sin \pi y \, \mathrm{d}y = \frac{1}{\pi} (1 - \cos \pi x)$. Hence $\vartheta_x = \alpha_{1-x} N(\alpha)_x = \frac{1}{\pi} (\sin \pi x) (1 - \cos \pi x) = \frac{1}{\pi} \sin \pi x - \frac{1}{2\pi} \sin 2\pi x$. The hypothesis for Proposition 8.2, that $\sqrt{\vartheta_x} \in \mathrm{L}^2[(0,1)]$ holds and so the α -weighted involutive walk is reversible with invariant distribution proportional to ϑ_x . Integrating over [0,1], the first summand gives $\frac{2}{\pi^2}$ and the

second 0, hence $\pi_x = \frac{\pi}{2}(\sin \pi x)(1 - \cos \pi x)$ is an invariant distribution. The kernel K in (8.1) satisfies

$$K(x,z)^2 = [x+z \ge 1] \frac{\sin \pi z \sin \pi x}{(1-\cos \pi x)(1-\cos \pi z)}.$$

Using $\sin w/(1-\cos w)=\cot \frac{w}{2}$, this simplifies to $K(x,z)^2=[x+z\geqslant 1]\cot \frac{\pi}{2}z\cot \frac{\pi}{2}x$. We have

$$\int_{\substack{(x,z)\in[0,1]^2\\x+z>1}} \cot\frac{\pi}{2}z\cot\frac{\pi}{2}z\cot\frac{\pi}{2}x\,\mathrm{d}z\,\mathrm{d}x = \int_0^1 \int_0^1 x\tan\frac{\pi}{2}xu\cot\frac{\pi}{2}x\,\mathrm{d}u\,\mathrm{d}x$$

by the substitution 1-z=xu. Since $x\tan\frac{\pi}{2}xu\cot\frac{\pi}{2}x\leqslant x\tan\frac{\pi}{2}x\cot\frac{\pi}{2}x=x$, the integrand is bounded. (In fact it is even continuous, except at (1,1).) Therefore $K\in L^2([0,1]^2)$ and the general theory in the previous subsection applies. In particular, we may apply Proposition 8.5 to conclude that π_x is the unique invariant distribution for the α -weighted involutive walk on [0,1], and that the process converges, in the Hilbert space sense, to the invariant distribution π_x .

We now find the spectra of L_P and R_P . In §4 and §5 we used righteigenvectors, corresponding to left-multiplication by the transition matrix, and again here it is most convenient to work with $L_P : \mathcal{H}_L \to \mathcal{H}_L$, defined as in (8.2) by

$$(L_P g)(x) = \pi \int_{1-x}^{1} \frac{\sin \pi z}{1 - \cos \pi x} g(z) dz,$$

where $\mathcal{H}_L = \{f(x)/\sqrt{\vartheta_x} : f \in L^2([0,1])\}$. For $n \in \mathbb{N}_0$, let V_n be the n-dimensional subspace of \mathcal{H}_L spanned by $1, \cos \pi x, \cos 2\pi x, \ldots, \cos \pi (n-1)x$ and let $V_\infty = \bigcup_{n=0}^\infty V_n$. The image of multiplication by $\sqrt{\pi_x}$ on $L^2([0,1])$ is dense in $L^2([0,1])$. (For example, it contains any step function supported on a closed subinterval of (0,1).) Since $\sqrt{2}, \sqrt{2}\cos \pi x, \sqrt{2}\cos 2\pi x, \ldots$ is an orthonormal Hilbert space basis of $L^2([0,1])$, the set $\sqrt{\pi_x}, \sqrt{\pi_x}\cos \pi x, \sqrt{\pi_x}\cos 2\pi x, \ldots$ is dense in $L^2([0,1])$. Hence, dividing by $\sqrt{\pi_x}$, we see that V_∞ is dense in \mathcal{H}_L .

Lemma 8.6. For any $d \in \mathbb{N}_0$ we have

$$L_P(\cos^d \pi x) \in \frac{(-1)^d}{d+1} \cos^d \pi x + V_d.$$

Proof. We have

$$L_P(\cos^d \pi x) = \frac{\pi}{1 - \cos \pi x} \int_{1-x}^1 \sin \pi z \cos^d \pi z \, dz$$
$$= \frac{\pi}{1 - \cos \pi x} \frac{-1}{\pi (d+1)} (\cos^{d+1} \pi z) \Big|_{1-x}^1$$
$$= \frac{-1}{d+1} \frac{(-1)^{d+1} - \cos^{d+1} \pi (1-x)}{1 - \cos \pi x}$$

$$= \frac{(-1)^d}{d+1} \frac{1 - \cos^{d+1} \pi x}{1 - \cos \pi x}$$
$$= (-1)^d \frac{1 + \cos \pi x + \dots + \cos^d \pi x}{d+1}$$

as required.

Since $\langle 1, \cos \pi x, \dots, \cos^d \pi x \rangle = \langle 1, \cos \pi x, \dots, \cos d \pi x \rangle$ for each d, it follows that L_P has eigenvalues $(-1)^d/(d+1)$ for $d \in \mathbb{N}_0$. Moreover, since V_{∞} is dense in \mathcal{H}_L , the corresponding eigenfunctions form a complete orthonormal basis for \mathcal{H}_L . The convergence of the involutive walk is determined by the second largest eigenvalue (in modulus), namely -1/2.

Remark 8.7. We remark that a similar method to Lemma 8.6 gives part of the spectrum of L_P whenever $\alpha_{1-x} = \alpha_x$ for all $x \in [0,1]$ and so $\alpha_{1-z}N(\alpha)_z^d$ is the derivative of $\frac{1}{d+1}N(\alpha)_z^{d+1}$. But one cannot expect in general that the eigenfunctions obtained in this way will be complete for the Hilbert space \mathcal{H}_L .

9. The polynomially-weighted involutive walk on the interval

Fix, throughout this section, $a, b \in \mathbb{N}_0$. Observe that the weight $\gamma^{(a,b)} =$ $\binom{y+a}{y}\binom{b+x-y}{x-y}$ defined in (1.2) is asymptotically proportional to $y^a(x-y)^b$ when y and x-y are both large. Let $\kappa^{(a,b)}$ be the real weight defined by $\kappa_{[y,x]}^{(a,b)} = y^a(x-y)^b$. In this section we use the theory from §8 to show that the $\kappa^{(a,b)}$ -weighted involutive walk, is, in a precise sense, the continuous limit of the $\gamma^{(a,b)}$ -involutive walk. In particular, it has a discrete spectrum and its eigenvalues are given by letting n tend to infinity in Theorem 1.5. We end with some final remarks comparing the discrete and continuous cases.

9.1. Polynomial weights. We begin with the analogues of Lemmas 5.1 and 5.2. Let $P(\kappa^{(a,b)})_{xz}$ be the probability density of a step from x to z. Part (i) below is a standard result related to the beta function which we prove to make the paper more self-contained.

Lemma 9.1. We have

(i)
$$\int_0^1 w^a (1-w)^b dw = (a+b+1)^{-1} {a+b \choose a}^{-1}$$
;

(i)
$$\int_0^1 w^a (1-w)^b dw = (a+b+1)^{-1} {a+b \choose a}^{-1};$$

(ii) $N(\kappa^{(a,b)})_x = x^{a+b+1} (a+b+1)^{-1} {a+b \choose b}^{-1};$

(iii)
$$P(\kappa^{(a,b)})_{xz} = [x+z \ge 1](a+b+1)\binom{a+b}{b} \frac{(1-z)^a(x+z-1)^b}{x^{a+b+1}}$$
 where $[x+z \ge 1]$ is an Iverson bracket.

Proof. By integrating $(w+(1-w)t)^c = \sum_{b=0}^c {c \choose b} w^{c-b} (1-w)^b t^b$ over $w \in [0,1]$ we get

$$\frac{1 - t^{c+1}}{(c+1)(1-t)} = \sum_{b=0}^{c} {c \choose b} t^b \int_0^1 w^{c-b} (1-w)^b \, \mathrm{d}y.$$

Since the left-hand side is $\frac{1}{c+1}(1+t+\cdots+t^c)$, (i) follows by setting c=a+b and comparing coefficients of t^b . Part (ii) follows from (i) using the change of variables xw=y to write $\int_0^x y^a(x-y)^b\,\mathrm{d}y=x^{a+b}\int_0^1 w^a(1-w)^bx\,\mathrm{d}w$. Now, starting at $x\in[0,1]$, we may step to $z\in[0,1]$ if and only if $1-z\leqslant x$; hence $P(\kappa^{(a,b)})_{xz}=[1-z\leqslant x]\kappa^{(a,b)}_{[1-z,x]}/N(\kappa^{(a,b)})_x$. Part (iii) now follows from the definition of $\kappa^{(a,b)}_{[1-z,x]}$ and (ii).

Set $\pi_x = C(1-x)^a x^{a+b+1}$ where $C = (2a+b+2)\binom{2a+b+1}{a}$ is the normalization factor given by Lemma 9.1(i).

Lemma 9.2. The $\kappa^{(a,b)}$ -interval involutive walk has unique invariant probability density function π_x and is reversible.

Proof. This is immediate from Proposition 8.2.

9.2. **Spectrum.** We now find the spectrum of the $\kappa^{(a,b)}$ -weighted involutive walks. The Hilbert space \mathcal{H}_L defined in §8.2 is as follows.

Definition 9.3. Let $\mathcal{H}_L = \{f(x)/\sqrt{\pi_x} : f \in L^2([0,1])\}$, with inner product $\langle , \rangle_{\mathcal{H}_L}$ defined by $\langle f, g \rangle = \int_0^1 (1-x)^a x^{a+b+1} f(x) g(x) dx$.

For $e \in \mathbb{N}$, let V_e be the subspace of \mathcal{H}_L of polynomials of degree strictly less than e, and let $V_{\infty} = \bigcup_{e \in \mathbb{N}} V_e$. Thus each V_e is a closed linear subspace of \mathcal{H}_L of dimension e and by a similar argument to §8.3, V_{∞} is a dense subspace of \mathcal{H}_L . By (8.2), the analogue of left-multiplication by the transition matrix is $L_P : \mathcal{H}_L \to \mathcal{H}_L$, defined by

$$(9.1) (L_P f)(x) = (a+b+1) \binom{a+b}{a} \int_{1-x}^1 \frac{(1-z)^a (x+z-1)^b}{x^{a+b+1}} f(z) dz.$$

We also need the analogous operator for left-multiplication by the transition matrix for down-steps,

(9.2)
$$(L_H f)(x) = (a+b+1) \binom{a+b}{a} \int_0^x \frac{y^a (x-y)^b}{x^{a+b+1}} f(y) \, \mathrm{d}y.$$

Changing variables from y to 1-z in the definition of $L_H f$ shows that $L_P = L_H J$, where $J: \mathcal{H}_L \to \mathcal{H}_L$ is the self-adjoint involution defined by (Jf)(x) = f(1-x), analogous to the matrices J(n) used earlier. Since J and L_P preserve \mathcal{H}_L , the same holds for L_H . Recall from Theorem 1.5 that the eigenvalues of the discrete $\gamma^{(a,b)}$ -weighted involutive walk on $\{0,\ldots,n-1\}$ are $(-1)^d \lambda_d^{(a,b)}$ for $0 \leq d < n$, where $\lambda_d^{(a,b)} = \binom{a+d}{d} / \binom{a+b+d+1}{d}$.

Lemma 9.4. Let $d \in \mathbb{N}_0$. We have $L_H(x^d) = \lambda_d^{(a,b)} x^d$.

Proof. By the substitution y = xw in (9.2) and Lemma 9.1(i) we get

$$L_H(x^d) = (a+b+1) \binom{a+b}{a} \int_0^x \frac{y^a (x-y)^b}{x^{a+b+1}} y^d \, dy$$
$$= (a+b+1) \binom{a+b}{a} x^d \int_0^1 w^{a+d} (1-w)^b \, dw$$

$$=\frac{(a+b+1)\binom{a+b}{a}}{(a+b+d+1)\binom{a+b+d}{a+d}}x^d.$$

Now use $\binom{a+b+d}{a+d}/\binom{a+b}{a}=\binom{a+b+d}{d}/\binom{a+d}{a}$ to see that the eigenvalue is as claimed. \Box

We now show that the analogue of Proposition 2.3(i) holds.

Lemma 9.5. Let $d \in \mathbb{N}$. Then

$$L_P(x^d) \in (-1)^d \lambda_d^{(a,b)} x^d + V_d.$$

Proof. Since $L_P = L_H J$ and $(L_H J)(x^d) = L_H ((1-x)^d) \in (-1)^d x^d + V_d$, this is immediate from Lemma 9.4.

We can now give a complete description of the spectrum and eigenfunctions for L_P . Let $g_0(x), g_1(x), g_2(x), \ldots$ be the functions obtained by Gram–Schmidt orthonormalization applied to $1, x, x^2, \ldots$, regarded as elements of the Hilbert space \mathcal{H}_L .

Theorem 9.6. The operator $L_P: \mathcal{H}_L \to \mathcal{H}_L$ is compact and self-adjoint. It has eigenvalues $(-1)^d \lambda_d^{(a,b)}$ for $d \in \mathbb{N}_0$. The normalized eigenfunction for $(-1)^d \lambda_d^{(a,b)}$ is g_d . The eigenfunctions g_0, g_1, \ldots are an orthonormal Hilbert space basis for \mathcal{H}_L and $L_P(f) = \sum_{d=0}^{\infty} (-1)^d \lambda_d^{(a,b)} \langle g_d, f \rangle g_d$ for each $f \in \mathcal{H}_L$. Proof. Let $f, g \in \mathcal{H}_L$. Let $D = (a+b+1)\binom{a+b}{g}$. Since

$$\langle f, L_{P}g \rangle_{H} = D \int_{0}^{1} (1-x)^{a} x^{a+b+1} f(x) \int_{1-x}^{1} \frac{(1-z)^{a} (x+z-1)^{a+b+1}}{x^{a+b+1}} g(z) dz$$
$$= D \iint_{\substack{(x,z) \in [0,1]^{2} \\ x+z \geqslant 1}} (1-x)^{a} (1-z)^{a} (x+z-1)^{a+b+1} f(x) g(z) dx dz$$

is symmetric with respect to swapping f and g and x and z, we have $\langle f, L_P g \rangle_{\mathcal{H}_L} = \langle L_P f, g \rangle_{\mathcal{H}_L}$. By induction on d, it follows from Lemma 9.5 that for each $d \in \mathbb{N}_0$, there is a unique polynomial $g \in V_{\infty}$ of degree d such that $\langle g, g \rangle_{\mathcal{H}_L} = 1$ and $L_P g = (-1)^d \lambda_d^{(a,b)} g$. Since the eigenfunctions for distinct eigenvalues of L_P are orthogonal, we have $g = g_d$. Since the g_d form a Hilbert space basis for \mathcal{H}_L , we have the claimed spectral decomposition of L_P . It is clear from (5.6) that $\lambda_d^{(a,b)} \to 0$ as $d \to \infty$. Therefore L_P is compact.

Dually, we defined in §8.2 the Hilbert space $\mathcal{H}_R = \{f(x)\sqrt{\pi_x} : f \in L^2([0,1])\}$, with inner product $\langle f,g\rangle_{\mathcal{H}_R} = \int_0^1 \frac{f(x)g(x)}{(1-x)^a x^{a+b+1}} \,\mathrm{d}x$. The map S which sends $f \in \mathcal{H}_L$ to $(1-x)^a x^{a+b+1} f(x) \in \mathcal{H}_R$ is an isometric isomorphism from \mathcal{H}_L to \mathcal{H}_R . Let $R_P : \mathcal{H}_R \to \mathcal{H}_R$ be the linear operator corresponding to right-multiplication by the transition matrix of the interval involutive walk, defined by

$$(R_P f)(x) = (a+b+1) {a+b \choose a} \int_{1-x}^1 \frac{(1-x)^a (x+z-1)^b}{z^{a+b+1}} f(z) dz.$$

Recall that $\pi_x = C(1-x)^a x^{a+b+1}$. Set $\widetilde{g}_d(x) = (1-x)^a x^{a+b+1} g_d$ for $d \in \mathbb{N}_0$.

Theorem 9.7. The operator $R_P: \mathcal{H}_R \to \mathcal{H}_R$ is compact and self-adjoint. It has eigenvalues $(-1)^d \lambda_d^{(a,b)}$ for $d \in \mathbb{N}_0$. The normalized eigenfunction for $(-1)^d \lambda_d^{(a,b)}$ is \widetilde{g}_d . The eigenfunctions $\widetilde{g}_0, \widetilde{g}_1, \ldots$ are an orthonormal Hilbert space basis for \mathcal{H}_R and $R_P(f) = \sum_{d=0}^{\infty} (-1)^d \lambda_d^{(a,b)} \langle \widetilde{g}_d, f \rangle \widetilde{g}_d$ for each $f \in \mathcal{H}_R$.

Proof. The operators R_P and L_P are conjugate by S, in that $R_PS = SL_P$. The result is now immediate from Theorem 9.6.

Hence, started at any probability distribution in \mathcal{H}_R , the $\kappa^{(a,b)}$ -weighted involutive walk converges, in the Hilbert space sense, to π_x . As in the discrete case, the rate of convergence controlled by the second largest eigenvalue in absolute value, namely (a+1)/(a+b+2).

Remark 9.8. As one might expect from the discrete case, all the results of this section extend to $a,b\in\mathbb{R}^{>-1}$, interpreting the factor $(a+b+1)^{-1}\binom{a+b}{b}^{-1}$ in the norm $N(\kappa^{(a,b)})$ as the value $\mathrm{B}(a+1,b+1)$ of the Beta integral. For instance, if $a=b=-\frac{1}{2}$ then $N(\kappa^{(-\frac{1}{2},-\frac{1}{2})})_x=\int_0^x \frac{\mathrm{d}y}{\sqrt{y(1-y)}}=\pi=\mathrm{B}(\frac{1}{2},\frac{1}{2})$.

9.3. Comparing the discrete and continuous.

Jacobi polynomials. For $f, g \in \mathcal{H}_L$ we have

$$\langle f, g \rangle_{\mathcal{H}_L} = \int_0^1 (1 - x)^a x^{a+b+1} f(x) g(x) \, \mathrm{d}x$$
$$= 2^{-(2a+b+1)} \int_{-1}^1 (1 - y)^a (1 + y)^{a+b+1} f(\frac{1+y}{2}) g(\frac{1+y}{2}) \, \mathrm{d}y.$$

It follows that the polynomials $g_d(\frac{1+y}{2})$ defined on [-1,1] are orthonormal with respect to the inner product adapted from \mathcal{H}_L , namely $\langle F,G\rangle = \int_{-1}^1 (1-y)^a (1+y)^{a+b+1} F(y) G(y) \, \mathrm{d}y$. The families of orthogonal polynomials for weight functions of the form $(1-y)^A (1+y)^B$ are known as the Jacobi polynomials. Thus $g_d(\frac{1+y}{2})$, is up to a scalar, the Jacobi polynomial $J^{(a,a+b+1)}(y)$ of degree d with parameters a and a+b+1.

For comparison, in the discrete setting of §5.4, for each fixed n we described the right-eigenvectors $w(0),\ldots,w(n-1)$ of the $n\times n$ matrix $P(\gamma^{(a,b)})$ as the Gram–Schmidt orthonormalization of the basis $v(0),\ldots,v(n-1)$ of \mathbb{R}^n of columns of Pascal's Triangle, with respect to the inner product $\langle v,w\rangle=\sum_{x=0}^{n-1}\pi_xv_xw_x$, where π_x is proportional to $\binom{n-x}{a}\binom{x+1}{a+b+1}$. Observe that the scaling factor π_x for v_xw_x is asymptotically proportional to $(n-x)^ax^{a+b+1}$ when x and n-x are both large. Thus, rescaling to the interval [0,1] by setting y=x/n and letting $n\to\infty$ we obtain $(1-y)^a(1+y)^{a+b+1}$, the same weight function as for the Jacobi polynomials. The eigenvectors w(d) from §5.4 are therefore the discrete analogues of the Jacobi polynomials. In particular, if $\widehat{w}^{(n)}$ is the eigenvector $w(d)\in\mathbb{R}^n$ of $P(\gamma^{(a,b)})$, scaled to have maximum entry (in absolute value) 1, and $\widehat{J}^{(a,a+b+1)}$

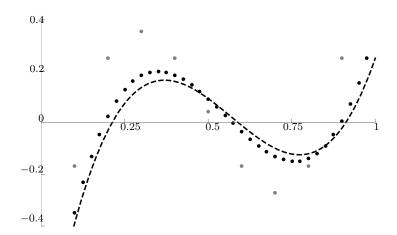


FIGURE 2. The right-eigenvector w(d) for $P(\gamma^{(a,b)})$ is the discrete analogue of the Jacobi polynomial with parameters a and a+b+1. The graph above shows the case a=0, b=0, rescaling to $[0,1]^2$, taking n=10 (grey dots) and n=40 (black dots) in the discrete case.

is the Jacobi polynomial $J_d^{(a,a+b+1)}$ with the same scaling imposed, then $\widehat{w}_{nx}^{(n)} \to \widehat{J}_d^{(a,a+b+1)}(2x-1)$ as $n \to \infty$ for each $x \in [0,1]$. In practice, the convergence is fast: see Figure 2 for an example. The results of §5.4 suggest these discrete analogues of the Jacobi polynomials deserve further combinatorial study.

Anti-diagonal conjugators and anti-diagonal eigenbasis action. We saw in §5.3 that the column v(d) of the binomial matrix B(n) is a right-eigenvector of the $n \times n$ down-step matrix $H(\gamma^{(a,b)})$ for the $\gamma^{(a,b)}$ -weighted involutive walk, and, in this basis, the transition matrix $P(\gamma^{(a,b)})$ is upper-triangular with diagonal entries $(-1)^d \lambda_d^{(a,b)}$. Theorem 2.9 shows that in this discrete case this behaviour (in the strong, global form) characterises the binomial transforms H^{λ} . Stated in a way we hope is intuitive, Proposition 2.8 says that the reason for this characterisation is that the unique basis of \mathbb{R}^{∞} projecting to bases of \mathbb{R}^n in which the operators J(n) for $n \in \mathbb{N}$ are all upper-triangular with alternating sign entries is given by the columns of the infinitely extended Pascal's triangle.

The trigonometric example in §8.3 and Theorem 9.6 show that this result has no immediate continuous analogue. Indeed, the operator $L_H: \mathcal{H}_L \to \mathcal{H}_L$ corresponding to left-multiplication by the down-step matrix in the trigonometric example in §8.3 is

$$L_H(g) = \pi \int_0^x \frac{\sin \pi z}{1 - \cos \pi x} g(z) dz.$$

As seen after Lemma 8.6, the operator L_P is represented by an (infinite) upper-triangular matrix in its action on these trigonometric eigenfunctions,

with diagonal entries $(-1)^d/(d+1)$. By Lemma 9.4 and Theorem 9.6 the analogous result holds replacing $\cos d\pi x$ with the monomials x^d for the right-multiplication operator for the $\kappa^{(a,b)}$ -weighted involutive walk. (The Jacobi function $g_d(2x-1)$ defined on [0,1] has degree d, so is a linear combination of $1, x, \ldots, x^d$.) Therefore both settings have the continuous analogue of the global anti-diagonal eigenvalue property, but the eigenfunctions are very different.

ACKNOWLEDGEMENTS

The authors thank an anonymous referee for helpful comments on an earlier version of this paper.

References

- [1] Stephen Boyd, Persi Diaconis, Pablo Parrilo, and Lin Xiao, Symmetry analysis of reversible Markov chains, Internet Math. 2 (2005), no. 1, 31–71.
- [2] F. P. Kelly, Reversibility and stochastic networks, Cambridge Mathematical Library, Cambridge University Press, Cambridge, 2011, Revised edition of the 1979 original with a new preface.
- [3] M. G. Kreĭn and M. A. Rutman, *Linear operators leaving invariant a cone in a Banach space*, Amer. Math. Soc. Translation **1950** (1950), no. 26, 128.
- [4] David A. Levin and Yuval Peres, Markov chains and mixing times, American Mathematical Society, Providence, RI, 2017, Second edition. With contributions by Elizabeth L. Wilmer, With a chapter on "Coupling from the past" by James G. Propp and David B. Wilson.
- [5] Eoghan McDowell, Tensor products of modular representations of $SL_2(\mathbb{F}_p)$ and a random walk on their indecomposable summands, arXiv:1904.13263 (April 2019), 30 pages.
- [6] Hiroyuki Ochiai, Makiko Sasada, Tomoyuki Shirai, and Takashi Tsuboi, Eigenvalue problem for some special class of anti-triangular matrices, arXiv:1403.6797 (March 2014), 23 pages.
- [7] Nicholas Young, An introduction to Hilbert space, Cambridge Mathematical Textbooks, Cambridge University Press, Cambridge, 1988.